# **Supervised Learning**

#### **Unsupervised learning:**

- ► To "extract structure" and postulate hypotheses about data generating process from observations x<sub>1</sub>,..., x<sub>n</sub>.
- Visualize, summarize and compress data.

We have seen how response or grouping variables are used to validate the usefulness of the extracted structure.

#### Supervised learning:

- ► In addition to the *n* observations of *X*, we also have a response variable  $Y \in \mathcal{Y}$ .
- Techniques for predicting Y given X.
  - Classification: discrete responses, e.g.  $\mathcal{Y} = \{+1, -1\}$  or  $\{1, \dots, K\}$ .
  - Regression: a numerical value is observed and  $\mathcal{Y} = \mathbb{R}$ .

Given training data  $(x_i, y_i)$ , i = 1, ..., n, the goal is to accurately predict the class or response *Y* on new observations of *X*.

## **Regression Example: Boston Housing**

The original data are 506 observations on 13 variables X; medv being the response variable Y.

crim	per capita crime rate by town		
zn	proportion of residential land zoned for lots		
	over 25,000 sq.ft		
indus	proportion of non-retail business acres per town		
chas	Charles River dummy variable (= 1 if tract bounds river;		
	0 otherwise)		
nox	nitric oxides concentration (parts per 10 million)		
rm	average number of rooms per dwelling		
age	proportion of owner-occupied units built prior to 1940		
dis	weighted distances to five Boston employment centers		
rad	index of accessibility to radial highways		
tax	full-value property-tax rate per USD 10,000		
ptratio	pupil-teacher ratio by town		
b	1000(B - 0.63)^2 where B is the proportion of blacks by to		
lstat	percentage of lower status of the population		
medv	median value of owner-occupied homes in USD 1000's		

## **Regression Example: Boston Housing**

> str(X)	
'data.frame': 506 obs. of 13 variables:	
\$ crim : num 0.00632 0.02731 0.02729 0.03237 0.06905 .	• • •
\$ zn : num 18 0 0 0 0 12.5 12.5 12.5 12.5	
\$ indus : num 2.31 7.07 7.07 2.18 2.18 2.18 7.87 7.87 7	7.87 7.87
\$ chas : int 0 0 0 0 0 0 0 0 0 0	
\$ nox : num 0.538 0.469 0.469 0.458 0.458 0.458 0.524	4 0.524 (
\$ rm : num 6.58 6.42 7.18 7.00 7.15	
\$ age : num 65.2 78.9 61.1 45.8 54.2 58.7 66.6 96.1 1	LOO 85.9
\$ dis : num 4.09 4.97 4.97 6.06 6.06	
\$ rad : int 1 2 2 3 3 3 5 5 5 5	
\$ tax : num 296 242 242 222 222 222 311 311 311 311 .	• • •
\$ ptratio: num 15.3 17.8 17.8 18.7 18.7 18.7 15.2 15.2 1	15.2 15.2
\$ black : num 397 397 393 395 397	
\$ lstat : num 4.98 9.14 4.03 2.94 5.33	

> str(Y)

num[1:506] 24 21.6 34.7 33.4 36.2 28.7 22.9 27.1 16.5 18.9 ...

Goal: predict median house price  $\hat{Y}(X)$ , given 13 predictor variables X of a new district.

## Classification Example: Lymphoma

We have gene expression measurements *X* of n = 62 patients for p = 4026 genes. For each patient, *Y* denotes one of two subtypes of cancer. Goal: predict cancer subtype  $\hat{Y}(X) \in \{0, 1\}$ , given gene expressions of a new patient.

> str(X)

'data.frame': 6	2 obs. of 4026 variables:
\$ Gene 1 : num	-0.344 -1.188 0.520 -0.748 -0.868
\$ Gene 2 : num	-0.953 -1.286 0.657 -1.328 -1.330
\$ Gene 3 : num	-0.776 -0.588 0.409 -0.991 -1.517
\$ Gene 4 : num	-0.474 -1.588 0.219 0.978 -1.604
\$ Gene 5 : num	-1.896 -1.960 -1.695 -0.348 -0.595
\$ Gene 6 : num	-2.075 -2.117 0.121 -0.800 0.651
\$ Gene 7 : num	-1.8755 -1.8187 0.3175 0.3873 0.0414
\$ Gene 8 : num	-1.539 -2.433 -0.337 -0.522 -0.668
\$ Gene 9 : num	-0.604 -0.710 -1.269 -0.832 0.458
\$ Gene 10 : num	-0.218 -0.487 -1.203 -0.919 -0.848
\$ Gene 11 : num	-0.340 1.164 1.023 1.133 -0.541
\$ Gene 12 : num	-0.531 0.488 -0.335 0.496 -0.358

> str(Y)

num [1:62] 0 0 0 1 0 0 1 0 0 0 ...

# **Decision Theory**

- Suppose we made a prediction  $\hat{Y} \in \mathcal{Y}$  based on observation of X.
- How good is the prediction? We can use a **loss function**  $L: \mathcal{Y} \times \mathcal{Y} \mapsto \mathbb{R}^+$  to formalize the quality of the prediction.
- Typical loss functions:
  - Misclassification loss (or 0-1 loss) for classification

$$L(Y, \hat{Y}) = \begin{cases} 0 & Y = \hat{Y} \\ 1 & Y \neq \hat{Y} \end{cases}.$$

Squared loss for regression

$$L(Y, \hat{Y}) = (Y - \hat{Y})^2.$$

► Alternative loss functions are often useful (later). For example, weighted misclassification error often appropriate. Or log-likelihood loss (sometimes shortened as log loss) L(Y, p̂) = -log p̂(Y), where p̂(k) is the estimated probability of class k ∈ Y.

## **Decision Theory**

For a given loss function L, the risk R of a learner is given by the expected loss

$$R(\hat{Y}) = \mathbb{E}(L(Y, \hat{Y}(X))),$$

where the expectation is with respect to the true (unknown) joint distribution (X, Y).

► The risk is unknown, but we can estimate it by the **empirical risk**:

$$R(\hat{Y}) \approx R_n(\hat{Y}) = \frac{1}{n} \sum_{i=1}^n L(y_i, \hat{Y}(x_i)).$$

- What is the optimal classifier if the joint distribution (X, Y) were known?
- ► The joint distribution *f* of *X* can be written as a mixture

$$f(X) = \sum_{k=1}^{K} f_k(X) \mathbb{P}(Y = k),$$

where, for  $k = 1, \ldots, K$ ,

- the prior probabilities over classes are  $P(Y = k) = \pi_k$
- and distributions of *X*, conditional on Y = k, is  $f_k(X)$ .
- The **Bayes classifier**  $\hat{Y}(X) \mapsto \{1, \ldots, K\}$  is the one with minimum risk:

$$R(\hat{Y}) = \mathbb{E}\left[L(Y, \hat{Y}(X))\right] = \mathbb{E}\left[\mathbb{E}[L(Y, \hat{Y}(x)|X = x]\right]$$
$$= \int_{\mathcal{X}} \mathbb{E}\left[L(Y, \hat{Y}(x))|X = x\right]f(x)dx$$

- ► The minimum risk attained by the Bayes classifier is called **Bayes risk**.
- Minimizing  $\mathbb{E}[L(Y, \hat{Y}(x)) | X = x]$  separately for each x suffices.

- Consider the situation of the 0-1 loss.
- The risk simplifies to:

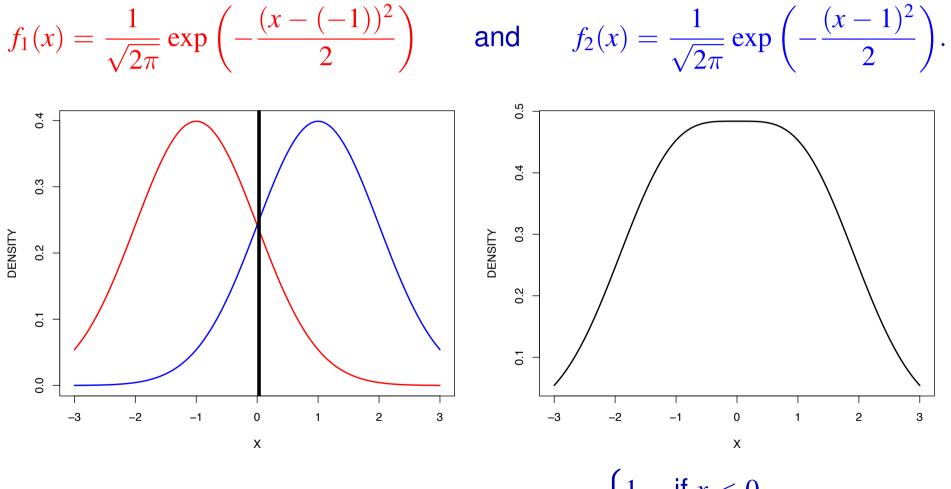
$$\mathbb{E}\left[L(Y,\hat{Y}(x))\big|X=x\right] = \sum_{k=1}^{K} L(k,\hat{Y}(x))\mathbb{P}(Y=k|X=x)$$
$$= 1 - \mathbb{P}(Y=\hat{Y}(x)|X=x)$$

The risk is minimized by choosing the class with the greatest posterior probability:

$$\hat{Y}(x) = \arg \max_{k=1,...,K} \mathbb{P}(Y = k | X = x) = \arg \max_{k=1,...,K} \frac{\pi_k f_k(x)}{\sum_{k=1}^K \pi_k f_k(x)} \\ = \arg \max_{k=1,...,K} \pi_k f_k(x).$$

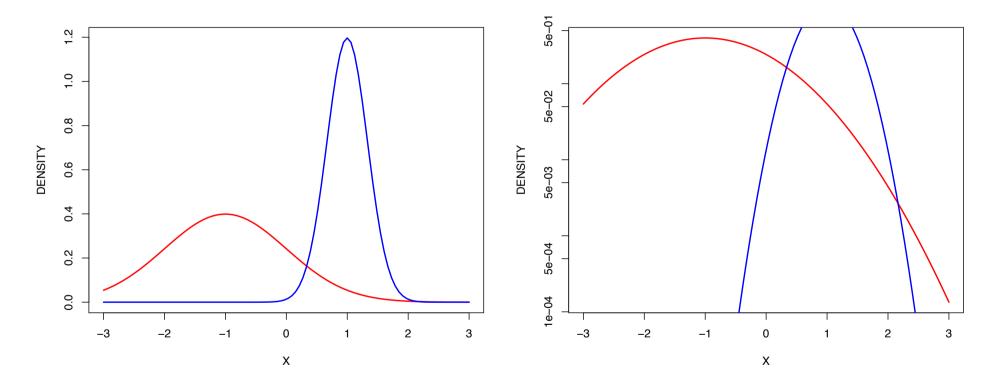
• The functions  $x \mapsto \pi_k f_k(x)$  are called **discriminant functions**. The function with maximum value determines the predicted class of *x*.

A simple two Gaussians example: Suppose  $X \sim \mathcal{N}(\mu_Y, 1)$ , where  $\mu_1 = -1$  and  $\mu_2 = 1$  and assume equal priors  $\pi_1 = \pi_2 = 1/2$ .



Optimal classification is  $\hat{Y}(x) = \underset{k=1,...,K}{\operatorname{arg\,max}} \pi_k f_k(x) = \begin{cases} 1 & \text{if } x < 0, \\ 2 & \text{if } x \ge 0. \end{cases}$ 

How do you classify a new observation x if now the standard deviation is still 1 for class 1 but 1/3 for class 2?



Looking at density in a log-scale, optimal classification is class 2 if and only if  $x \in [-0.39, 2.15]$ .

# **Plug-in Classification**

The Bayes Classifier chooses the class with the greatest posterior probability

 $\hat{Y}(x) = \underset{k=1,...,K}{\operatorname{arg max}} \pi_k f_k(x).$ 

- Unfortunately, we usually know neither the conditional class probabilities nor the prior probabilities.
- We can estimate the joint distribution with:
  - estimates  $\hat{\pi}_k$  for  $\pi_k$  and  $k = 1, \ldots, K$  and
  - estimates  $\hat{f}_k(x)$  of conditional class densities,
- ► The **plug-in classifiers** chooses the class

$$\hat{Y}(x) = \operatorname*{arg\,max}_{k=1,\ldots,K} \hat{\pi}_k \hat{f}_k(x).$$

Linear Discriminant Analysis will be an example of plug-in classification.

## Linear Discriminant Analysis

- LDA is the most well-known and simplest example of plug-in classification.
- Assume a multivariate Normal form for  $f_k(x)$  for each class k:

 $X|Y = k \sim \mathcal{N}(\mu_k, \Sigma),$ 

- each class can have a **different mean**  $\mu_k$
- but all classes share the same covariance  $\Sigma$ .
- ► For an observation *x*,

$$\log \mathbb{P}(Y = k | X = x) = \kappa + \log \pi_k f_k(x)$$
$$= \kappa + \log \pi_k - \frac{1}{2} (x - \mu_k)^\top \Sigma^{-1} (x - \mu_k)$$

The quantity  $(x - \mu_k)^\top \Sigma^{-1} (x - \mu_k)$  is the square of the **Mahalanobis distance**. It gives the distance between *x* and  $\mu_k$  in the metric given by  $\Sigma$ .

• If  $\Sigma = I_p$  and  $\pi_k = \frac{1}{K}$ ,  $\hat{Y}(x)$  simply chooses the class k with the nearest (in the Euclidean sense) mean.

#### **Linear Discriminant Analysis**

• Expanding the **discriminant**  $(x - \mu_k)^\top \Sigma^{-1} (x - \mu_k)$ ,

$$\log \mathbb{P}(Y = k|x) = \kappa + \log(\pi_k) - \frac{1}{2} \left( \mu_k^\top \Sigma^{-1} \mu_k - 2\mu_k^\top \Sigma^{-1} x + x^\top \Sigma^{-1} x \right)$$
$$= \kappa + \log(\pi_k) - \frac{1}{2} \mu_k^\top \Sigma^{-1} \mu_k + \mu_k^\top \Sigma^{-1} x$$

• Setting  $a_k = \log(\pi_k) - \frac{1}{2}\mu_k^\top \Sigma^{-1}\mu_k$  and  $b_k = \Sigma^{-1}\mu_k$ , we obtain

$$\log \mathbb{P}(Y = k | X = x) = \kappa + a_k + b_k^\top x$$

i.e. a linear discriminant function.

Consider choosing class k over k':

$$a_k + b_k^\top x > a_{k'} + b_{k'}^\top x \qquad \Leftrightarrow \qquad a_\star + b_\star^\top x > 0$$

where  $a_{\star} = a_k - a_{k'}$  and  $b_{\star} = b_k - b_{k'}$ .

- The Bayes classifier partitions X into regions with the same class predictions via separating hyperplanes.
- The Bayes classifier under these assumptions is more commonly known as the LDA classifier.

#### **Parameter Estimation**

- The final piece of the puzzle is to estimate the parameters of the LDA model.
- We can achieve this by maximum likelihood.
- EM algorithm is not needed here since the class variables y<sub>i</sub> are observed.
- Let  $n_k = \#\{j : y_j = k\}$  be the number of observations in class k.

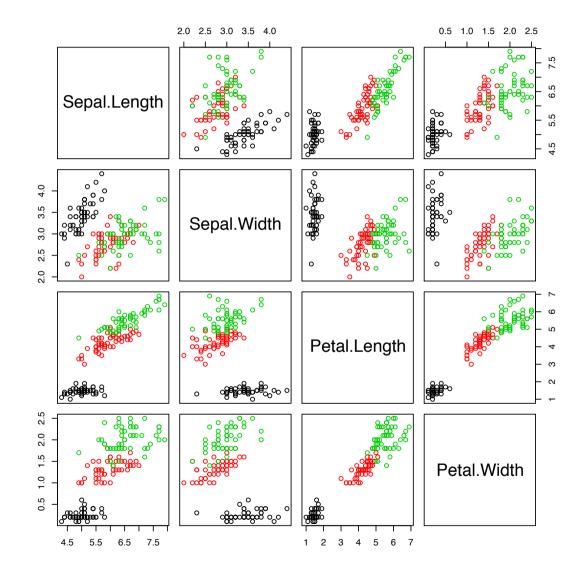
$$\ell(\pi, (\mu_k), \Sigma) = \kappa + \sum_{k=1}^{K} \sum_{j: y_j = k} \log \pi_k - \frac{1}{2} \left( \log |\Sigma| + (x_j - \mu_k)^\top \Sigma^{-1} (x_j - \mu_k) \right)$$

Then:

$$\hat{\pi}_k = \frac{n_k}{n} \qquad \qquad \hat{\mu}_k = \frac{1}{n_k} \sum_{j:y_j=k} x_j$$
$$\hat{\Sigma} = \frac{1}{n} \sum_{k=1}^K \sum_{j:y_j=k} (x_j - \hat{\mu}_k) (x_j - \hat{\mu}_k)^\top$$

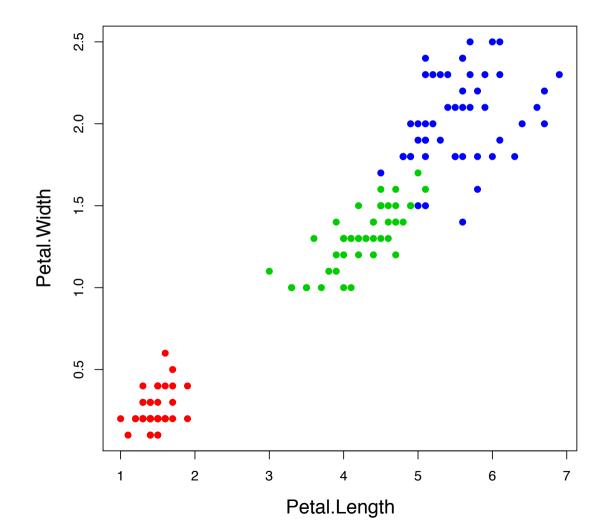
Note: the ML estimate of  $\Sigma$  is not unbiased. For an unbiased estimate we need to divide by n - K.

```
library(MASS)
data(iris)
##save class labels
ct <- rep(1:3,each=50)
##pairwise plot
pairs(iris[,1:4],col=ct)</pre>
```



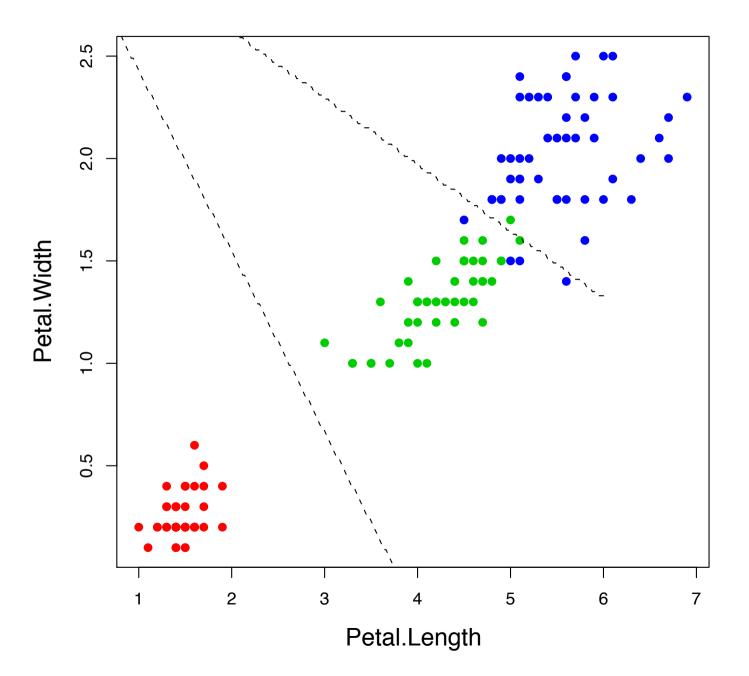
Just focus on two predictor variables.

```
iris.data <- iris[,3:4]
plot(iris.data,col=ct+1,pch=20,cex=1.5,cex.lab=1.4)</pre>
```



#### Computing and plotting the LDA boundaries.

```
##fit LDA
iris.lda <- lda(x=iris.data,grouping=ct)
##create a grid for our plotting surface
x <- seq(-6,6,0.02)
y <- seq(-4,4,0.02)
z <- as.matrix(expand.grid(x,y),0)
m <- length(x)
n <- length(y)
##classes are 1,2 and 3, so set contours at 1.5 and 2.5
```

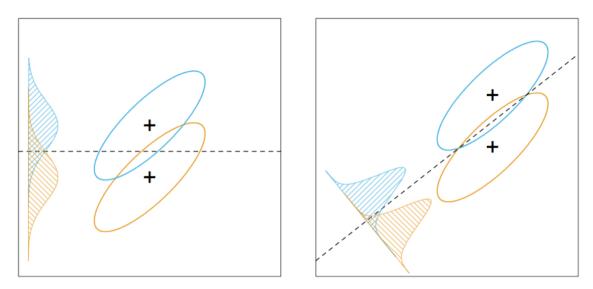


## Fisher's Linear Discriminant Analysis

- ▶ In LDA, data vectors are classified based on Mahalanobis distance from cluster means, which lie on a K 1 affine subspace.
- In measuring these distances, directions orthogonal<sup>5</sup> to the subspace can be ignored.
- Projecting data vectors onto the subspace can be viewed as a dimensionality reduction technique that preserves discriminative information about (y<sub>i</sub>)<sup>n</sup><sub>i=1</sub>.
- As with PCA, we can visualize the structure in the data by choosing an appropriate basis for the subspace and projecting data onto it.
- Choose a basis by finding directions that are separate classes best.

<sup>&</sup>lt;sup>5</sup>Orthogonality defined in terms of the inner product corresponding to Mahalanobis distance:  $\langle x, y \rangle = x \Sigma^{-1} y$ .

#### Fisher's Linear Discriminant Analysis



Find a direction  $v \in \mathbb{R}^p$  to maximize the variance ratio

$$\frac{v^{\top} B v}{v^{\top} \Sigma v}$$

where

$$\Sigma = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \mu_{y_i}) (x_i - \mu_{y_i})^{\top}$$
$$B = \frac{1}{n-1} \sum_{k=1}^{K} n_k (\mu_{y_i} - \bar{x}) (\mu_{y_i} - \bar{x}))^{\top}$$
$$B \text{ has rank at most } K - 1.$$

(within class covariance)(between class covariance)

# **Discriminant Coordinates**

• To solve for the optimal v, we first reparameterize it as  $u = \sum_{i=1}^{\frac{1}{2}} v$ .

$$\frac{v^{\top}Bv}{v^{\top}\Sigma v} = \frac{u^{\top}(\Sigma^{-\frac{1}{2}})^{\top}B\Sigma^{-\frac{1}{2}}u}{u^{\top}u} = \frac{u^{\top}B^{*}u}{u^{\top}u}$$

where  $B^* = (\Sigma^{-\frac{1}{2}})^{\top} B \Sigma^{-\frac{1}{2}}$ .

- The maximization over u is achieved by the first eigenvector  $u_1$  of  $B^*$ .
- We also look at the remaining eigenvectors  $u_l$  associated to the non-zero eigenvalues and defined the **discriminant coordinates** as  $v_l = \sum_{l=1}^{-1} u_l$ .
- The  $v_l$ 's span exactly the affine subspace spanned by  $(\Sigma^{-1}\mu_k)_{k=1}^K$  (these vectors are given as the "linear discriminants" in the R-function 1da).

library(MASS)
data(crabs)

```
## numeric and text class labels
ct <- as.numeric(crabs[,1])-1+2*(as.numeric(crabs[,2])-1)</pre>
```

## Projection on Fisher's linear discriminant directions
print(cb.lda <- lda(log(crabs[,4:8]),ct))</pre>

> > > > > > > > > > > > Call:
lda(log(crabs[, 4:8]), ct)

Prior probabilities of groups: 0 1 2 3 0.25 0.25 0.25 0.25

Group means:

FLRWCLCWBD02.5649852.4751743.3126853.4623272.44135112.8524552.6838313.5293703.6495552.73327322.6727242.4437743.4379683.5780772.56080632.7878852.4899213.4904313.5894262.701580

Coefficients of linear discriminants:

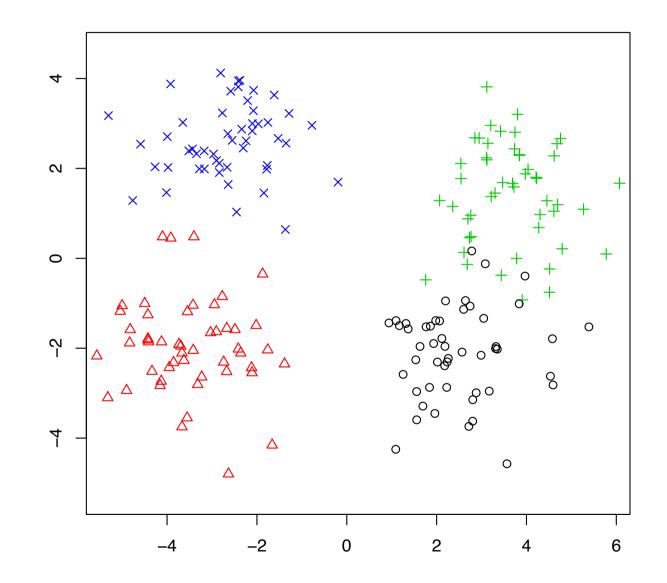
	LD1	LD2	LD3
FL	-31.217207	-2.851488	25.719750
RW	-9.485303	-24.652581	-6.067361
CL	-9.822169	38.578804	-31.679288
CW	65.950295	-21.375951	30.600428
BD	-17.998493	6.002432	-14.541487

 Proportion of trace:

 LD1
 LD2
 LD3

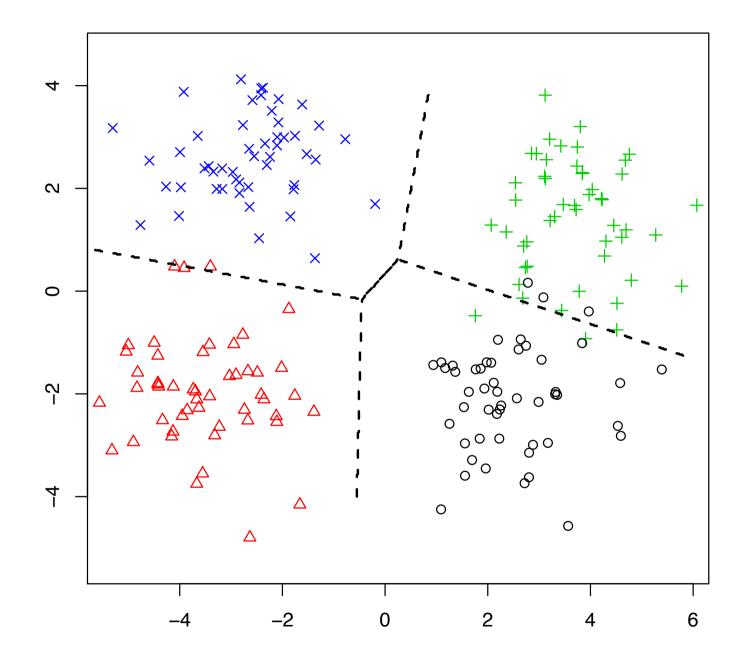
 0.6891
 0.3018
 0.0091

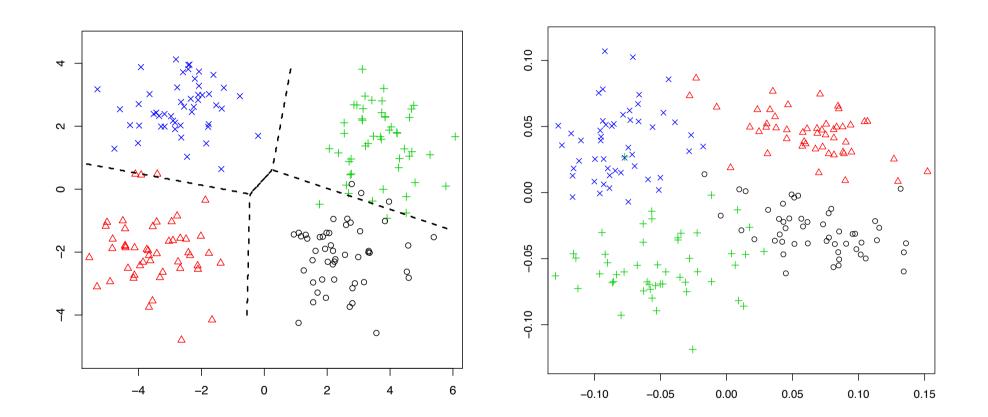
cb.ldp <- predict(cb.lda)
eqscplot(cb.ldp\$x,pch=ct+1,col=ct+1)</pre>



```
## display the decision boundaries
## take a lattice of points in LD-space
x <- seq(-6,6,0.02)
y <- seq(-4,4,0.02)
z <- as.matrix(expand.grid(x,y,0))
m <- length(x)
n <- length(y)</pre>
```

```
## predict onto the grid
cb.ldap <- lda(cb.ldp$x,ct)
cb.ldpp <- predict(cb.ldap,z)$class</pre>
```





LDA separates the groups better.

## Naïve Bayes

- Assume we are interested in classifying documents; e.g. scientific articles or emails.
- A basic but standard model for text classification consists of considering a pre-specified dictionary of *p* words (including say physics, calculus.... or dollars, sex etc.) and summarizing each document *i* by a binary vector *x<sub>i</sub>* where

$$x_{ij} = \begin{cases} 1 & \text{if word } j \text{ is present in document} \\ 0 & \text{otherwise.} \end{cases}$$

• To implement a probabilistic classifier, we need to model  $f_k(x|\phi_k)$  for each class k = 1, ..., K.

## Naïve Bayes

• A Naïve Bayes approach ignores feature correlations and assumes  $f_k(x) = f(x|\phi_k)$  where

$$f_k(x_i) = f(x_i | \phi_k) = \prod_{j=1}^p (\phi_{kj})^{x_{ij}} (1 - \phi_{kj})^{1 - x_{ij}}$$

Given dataset, the MLE is easily obtained

$$\hat{\pi}_k = \frac{n_k}{n} \qquad \qquad \hat{\phi}_{kj} = \frac{\sum_{i:y_i=k} x_{ij}}{n_k}$$

• One problem: if word *j* did not appear in documents labelled as class *k* then  $\hat{\phi}_{kj} = 0$  and

 $\mathbb{P}(Y = k | X = x \text{ with } j \text{ th entry equal to } 1) = 0$ 

i.e. we will never attribute a new document containing word j to class k.

This problem is called **overfitting**, and is a major concern in modelling high-dimensional datasets common in machine learning.