Graphical and Log-Linear Models

MSc Further Statistical Methods, Lecture 2 Hilary Term 2006

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Conditional independence

Random variables X and Y are *conditionally independent* given the random variable Z if

$$\mathcal{L}(X \mid Y, Z) = \mathcal{L}(X \mid Z).$$

We then write $X \perp\!\!\!\perp Y \mid Z$

Intuitively:

Knowing Z renders Y irrelevant for predicting X.

Factorisation of probabilities:

$$X \perp \!\!\!\perp Y \mid Z \quad \Longleftrightarrow \quad p_{xyz}p_{++z} = p_{x+z}p_{+y,z} \\ \iff \quad \exists a, b : p_{xyz} = a_{yz}b_{yz}.$$

Fundamental properties

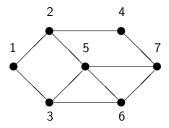
For any random variables X, Y, Z, and W it holds

- (C1) if $X \perp\!\!\!\perp Y \mid Z$ then $Y \perp\!\!\!\perp X \mid Z$;
- (C2) if $X \perp\!\!\!\perp Y \mid Z$ and U = g(Y), then $X \perp\!\!\!\perp U \mid Z$;
- (C3) if $X \perp\!\!\!\perp Y \mid Z$ and U = g(Y), then $X \perp\!\!\!\perp Y \mid (Z, U)$;
- (C4) if $X \perp\!\!\!\perp Y \mid Z$ and $X \perp\!\!\!\perp W \mid (Y, Z)$, then $X \perp\!\!\!\perp (Y, W) \mid Z$;

If all joint probabilities $p_{xyzw} \ {\rm are \ strictly \ positive \ also}$

(C5) if
$$X \perp\!\!\!\perp Y \mid (Z, W)$$
 and $X \perp\!\!\!\perp Z \mid (Y, W)$ then $X \perp\!\!\!\perp (Y, Z) \mid W$.

Graphical models



For several variables, complex systems of conditional independence can be described by undirected graphs.

A set of variables A is conditionally independent of set B, given the values of a set of variables C if C separates A from B.

For example in picture above

$$1 \perp\!\!\!\perp \{4,7\} \,|\, \{2,3\}, \qquad \{1,2\} \perp\!\!\!\perp 7 \,|\, \{4,5,6\}.$$

Algebraically the picture represents the fact that the joint probability of all variables factorizes into terms that only depends on *cliques* of the graph. In the pictures:

$$p_{ijklmno} = a_{ij}b_{ik}c_{jm}c_{jl}c_{kmn}c_{lo}c_{mno}.$$

Alternatively, the graph can be interpreted by saying that for each missing edge, there is a conditional independence associated:

$$I \not\sim J \implies I \perp \!\!\!\perp J \mid \mathsf{remaining variables}.$$

Formal Markov properties

Formally we say for a given graph $\ensuremath{\mathcal{G}}$ that a distribution obeys

(P) the pairwise Markov property if

 $\alpha \not \sim \beta \implies \alpha \amalg \beta \, | \, V \setminus \{ \alpha, \beta \},$

i.e. if all non-neighbours are conditionally independent given the remaining;

(L) the local Markov property if

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\forall \alpha \in V : \alpha \perp V \setminus \operatorname{cl}(\alpha) \mid \operatorname{bd}(\alpha),
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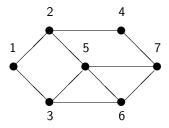
i.e. every variable is conditionally independent of the remaining given its neighbours;

(G) the global Markov property if

S separates A from B implies $A \perp\!\!\!\perp B \mid S$.

These are all equivalent if probabilities are strictly positive.

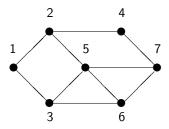
Pairwise Markov property



Any non-adjacent pair of random variables are conditionally independent given the remaning.

For example, $1 \perp 5 \mid \{2, 3, 4, 6, 7\}$ and $4 \perp 6 \mid \{1, 2, 3, 5, 7\}$.

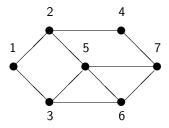
Local Markov property



Every variable is conditionally independent of the remaining, given its neighbours.

For example, $5 \perp\!\!\!\perp \{1,4\} \mid \! \{2,3,6,7\}$ and $7 \perp\!\!\!\perp \{1,2,3\} \mid \! \{4,5,6\}.$

Global Markov property



To find conditional independence relations, one should look for separating sets, such as $\{2,3\}$, $\{4,5,6\}$, or $\{2,5,6\}$ For example, it follows that $1 \perp 1 \mid \{2,5,6\}$ and

 $2 \perp\!\!\!\perp 6 \,|\, \{3, 4, 5\}.$

Log-linear models

 $\mathcal{A} = \{a_1, \ldots, a_K\}$ denotes a set of (pairwise incomparable) subsets of $a_i \subseteq V$.

A probability distribution p (or function) factorizes w.r.t. A if it can be written as a product of terms where each only depend on variables in the same subset of A.

The set of distributions which factorize w.r.t. \mathcal{A} is the log-linear model generated by \mathcal{A} .

 ${\cal A}$ is the generating class of the log–linear model.

It can be shown that a positive probability distribution is Markov w.r.t. a graph if and only if it factorizes as above with a_i being *complete sets*, i.e. sets where all elements are mutual neighbours. If the distribution factorizes without being everywhere positive, it will also satisfy all the Markov properties, but not the other way around.

Formally, we define the graphical model with graph G = (V, E) to be the log-linear model with $\mathcal{A} = \mathcal{C}$, where \mathcal{C} are the *cliques* (i.e. maximal complete subsets) of the graph.

Example

Consider a three way contingency table, where e.g. m_{ijk} denotes the mean of the counts N_{ijk} in the cell (i, j, k) which has then been expanded as e.g.

$$\log m_{ijk} = \alpha_i + \beta_j + \gamma_k \tag{1}$$

or

$$\log m_{ijk} = \alpha_{ij} + \beta_{jk} \tag{2}$$

or

$$\log m_{ijk} = \alpha_{ij} + \beta_{jk} + \gamma_{ik}, \qquad (3)$$

or (with redundancy)

$$\log m_{ijk} = \gamma + \delta_i + \phi_j + \eta_k + \alpha_{ij} + \beta_{jk} + \gamma_{ik}$$

The additive terms in the expansion are known as interaction terms of order |a| - 1 or |a|-factor interactions.

Interaction terms of 0th order are called *main effects*.

Dependence graph

Any joint probability distribution of has a dependence graph G = G(P) = (V, E(P)).

This is defined by letting $\alpha \not\sim \beta$ in G(P) exactly when

 $\alpha \, \bot\!\!\!\bot \, \beta \, | \, V \setminus \{ \alpha, \beta \}.$

X will then satisfy the pairwise Markov w.r.t. ${\cal G}(P)$ and the other Markov properties as well, in the case of positive probabilities.

Dependence graph of log-linear model

For any generating class \mathcal{A} we can construct the dependence graph of the corresponding log-linear model.

This is determined by the relation

$$\alpha \sim \beta \iff \exists a \in \mathcal{A} : \alpha, \beta \in a.$$

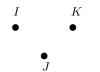
They are then also global, local, and pairwise Markov w.r.t. $G(\mathcal{A}).$

This is by default the graph displayed in MIM.

Independence

The log-linear model specified by (1) is known as the *main effects model*.

It has generating class consisting of singletons only $\mathcal{A}=\{\{I\},\{J\},\{K\}\}.$ It has dependence graph



Thus it corresponds to *complete independence*.

Conditional independence

The log-linear model specified by (2) has no interaction between I and K.

It has generating class $\mathcal{A} = \{\{I,J\},\{J,K\}\}$ and dependence graph



Thus it corresponds to the *conditional independence* $I \perp\!\!\!\perp K \mid J$.

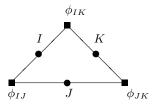
No interaction of second order

The log-linear model specified by (3) has no second-order interaction. It has generating class $\mathcal{A} = \{\{I, J\}, \{J, K\}, \{I, K\}\} \text{ and its dependence graph}$



is the complete graph. Thus it has no conditional independence interpretation.

Interaction graphs



The interaction graph of A is the graph with vertices $V \cup A$ and edges define by

$$\alpha \sim a \iff \alpha \in a.$$

Using this graph all log-linear models admit a simple visual representation. Can be requested in MIM.

Likelihood function

The likelihood function for an unknown p can be expressed as

$$L(p) = \prod_{\nu=1}^{n} p(x^{\nu}) = \prod_{x \in \mathcal{X}} p(x)^{n(x)}.$$

In contingency table form the data follow a multinomial distribution

$$P\{N(x) = n(x), x \in \mathcal{X}\} = \frac{n!}{\prod_{x \in X} n(x)!} \prod_{x \in \mathcal{X}} p(x)^{n(x)}$$

but this only affects the likelihood function by a constant factor.

It can be shown that in log-linear models, *the likelihood function has at most one maximum*. When zero-values are allowed, it *always* has one.

MIM uses an algorithm for fitting known as *Iterative Proportional Fitting* which, if properly implemented, also works in the case where probabilities are allowed to be zero (sparse tables).

Also implemented e.g. in R in loglin with front end loglm in MASS.

An alternative is to "pretend" that counts are independent and Poisson distributed and use glm. However, the algorithm used there does *not* work when estimated cell probabilities are zero.