

YEP VII "Probability, random trees and algorithms" 8th-12th March 2010

Scaling limits for random trees and graphs

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INTRODUCTION

A taste of what's to come

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Let $\mathbb{T}_{[n]}$ be the set of unordered trees on n vertices labelled by $[n] := \{1, 2, \dots, n\}.$

For example, $\mathbb{T}_{[3]}$ consists of the trees



Unordered trees

Note that unordered means that these trees are all the same:



but this one is different:



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Uniform random trees

Cayley's formula tells us that $|\mathbb{T}_{[n]}| = n^{n-2}$.

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What happens as *n* grows?

- 1. Fix $n \ge 2$.
- 2. Start from the vertex labelled 1.

3. For $2 \le i \le n$, connect vertex *i* to vertex V_i such that

$$V_i = egin{cases} j & ext{with probability } 1/n, \ 1 \leq j \leq i-2, \ i-1 & ext{with probability } 1-(i-2)/n. \end{cases}$$

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4. Take a uniform random permutation of the labels.

[See Nicolas Broutin's lecture.]

Consider n = 10.

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 $V_2 = 1$ with probability 1





$$V_3 = egin{cases} 1 & ext{with probability 1/10} \ 2 & ext{with probability 9/10} \end{cases}$$



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$$V_4 = egin{cases} j & ext{with probability 1/10, } 1 \leq j \leq 2 \ 3 & ext{with probability 8/10} \end{cases}$$



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$$V_5 = egin{cases} j & ext{with probability 1/10, } 1 \leq j \leq 3 \ 4 & ext{with probability 7/10} \end{cases}$$



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$$V_6 = egin{cases} j & ext{with probability 1/10, } 1 \leq j \leq 4 \ 5 & ext{with probability 6/10} \end{cases}$$



$$V_7 = egin{cases} j & ext{with probability 1/10, } 1 \leq j \leq 5 \ 6 & ext{with probability 5/10} \end{cases}$$



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$$V_8 = egin{cases} j & ext{with probability 1/10, } 1 \leq j \leq 6 \ 7 & ext{with probability 4/10} \end{cases}$$



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$$V_9 = egin{cases} j & ext{with probability 1/10, } 1 \leq j \leq 7 \\ 8 & ext{with probability 3/10} \end{cases}$$



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$$V_{10} = egin{cases} j & ext{with probability 1/10, } 1 \leq j \leq 8 \ 9 & ext{with probability 2/10} \end{cases}$$



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Permute.



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Typical distances

Consider the tree before we permute. Let $J_n = \inf\{i \ge 1 : V_{i+1} \neq i\}$. We can use J_n to give us an idea of typical distances in the tree.

In our example, $J_{10} = 4$:



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Typical distances

Proposition

$n^{-1/2}J_n$ converges in distribution as $n \to \infty$.



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Imagine now that edges in the tree have length 1. This result suggests that rescaling edge-lengths by $n^{-1/2}$ will give some sort of limit for the whole tree. The limiting version of the algorithm is as follows.

Take an inhomogeneous Poisson process on \mathbb{R}^+ of intensity *t* at *t*.



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Consider the line-segments $[0, C_1), [C_1, C_2), \ldots$

Take an inhomogeneous Poisson process on \mathbb{R}^+ of intensity *t* at *t*.



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Consider the line-segments $[0, C_1), [C_1, C_2), \ldots$

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For $i \ge 2$, attach $[C_{i-1}, C_i)$ at a random point chosen uniformly over the existing tree.





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Take the closure of the union of all the branches.

This procedure gives (a rather informally expressed) definition of Aldous' Brownian continuum random tree (CRT).

The Brownian continuum random tree



[Picture by Grégory Miermont]

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DISCRETE TREES

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Based in large part on **Random trees and applications** by Jean-François Le Gall.

It turns out to be more natural to work with rooted, ordered trees (also called plane trees).

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We will use the Ulam-Harris labelling. Let $\mathbb{N} = \{1, 2, 3, \ldots\}$ and

$$\mathcal{U}=\bigcup_{n=0}^{\infty}\mathbb{N}^{n},$$

where $\mathbb{N}^0 = \{\emptyset\}$.

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$$\mathcal{U}=\bigcup_{n=0}^{\infty}\mathbb{N}^{n},$$

where $\mathbb{N}^0 = \{\emptyset\}$. An element $u \in \mathcal{U}$ is a sequence $u = (u^1, u^2, \dots, u^n)$ of natural numbers representing a point in an infinitary tree:



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Thus the label of a vertex indicates its genealogy.

Write |u| = n for the generation of u.

u has parent $p(u) = (u^1, u^2, \dots, u^{n-1})$.

u has children u1, u2, ... where, in general, $uv = (u^1, u^2, ..., u^n, v^1, v^2, ..., v^m)$ is the concatenation of sequences $u = (u^1, u^2, ..., u^n)$ and $v = (v^1, v^2, ..., v^m)$.

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We root the tree at \emptyset .

A (finite) rooted, ordered tree \mathbf{t} is a finite subset of \mathcal{U} such that

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- ▶ for all $u \in \mathbf{t}$ such that $u \neq \emptyset$, $p(u) \in \mathbf{t}$
- ▶ for all $u \in \mathbf{t}$, there exists $k(u) \in \mathbb{Z}_+$ such that for $j \in \mathbb{N}$, $uj \in \mathbf{t}$ iff $1 \le j \le k(u)$.

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k(u) is the number of children of u in **t**.

Write #(t) for the size (number of vertices) of t and note that

$$\#(\mathbf{t}) = 1 + \sum_{u \in \mathbf{t}} k(u).$$

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k(u) is the number of children of u in **t**.

Write #(t) for the size (number of vertices) of t and note that

$$\#(\mathbf{t}) = 1 + \sum_{u \in \mathbf{t}} k(u).$$

Write **T** for the set of all rooted ordered trees.

Two ways of encoding a tree

Consider a rooted ordered tree $\mathbf{t} \in \mathbf{T}$.

It will be convenient to encode this tree in terms of discrete functions which are easier to manipulate.

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It will be convenient to encode this tree in terms of discrete functions which are easier to manipulate.

We will do this is two different ways:

- the height function
- the depth-first walk.

Suppose that **t** has *n* vertices. Let them be $v_0, v_1, \ldots, v_{n-1}$, listed in lexicographical order.

Suppose that **t** has *n* vertices. Let them be $v_0, v_1, \ldots, v_{n-1}$, listed in lexicographical order.

Then the height function is defined by

$$H(k) = |v_k|, \quad 0 \le k \le n-1.$$

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We can recover the tree from its height function (after a little thought!).

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Recall that k(v) is the number of children of v, and that $v_0, v_1, \ldots, v_{n-1}$ is a list of the vertices of **t** in lexicographical order.

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Recall that k(v) is the number of children of v, and that $v_0, v_1, \ldots, v_{n-1}$ is a list of the vertices of **t** in lexicographical order.

Define

$$X(0) = 0,$$

 $X(i) = \sum_{j=0}^{i-1} (k(v_j) - 1), ext{ for } 1 \leq i \leq n$

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Recall that k(v) is the number of children of v, and that $v_0, v_1, \ldots, v_{n-1}$ is a list of the vertices of **t** in lexicographical order.

Define

$$X(0) = 0,$$

 $X(i) = \sum_{j=0}^{i-1} (k(v_j) - 1), ext{ for } 1 \leq i \leq n.$

In other words,

$$X(i+1) = X(i) + k(v_i) - 1, \quad 0 \le i \le n-1.$$

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Step 0



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Current: Ø Alive: none Dead: none

Step 1



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Current: 1 Alive: none Dead: ∅

Step 2



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Current: 1 1 Alive: 1 2 Dead: \emptyset , 1

Step 3



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Current: 1 1 1 Alive: 1 1 2, 1 2 Dead: Ø, 1, 1 1

Step 4



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Current: 1 1 2 Alive: 1 2 Dead: Ø, 1, 1 1, 1 1 1

Step 5



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Current: 1 2 Alive: none Dead: Ø, 1, 1 1, 1 1 1, 1 1 2

Step 6



Current: 1 2 1 Alive: none Dead: Ø, 1, 1 1, 1 1 1, 1 1 2, 1 2

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Step 7



Dead: Ø, 1, 1 1, 1 1 1, 1 1 2, 1 2, 1 2 1



It is less easy to see that the depth-first walk also encodes the tree.

Proposition

For
$$0 \le i \le n - 1$$
,
 $H(i) = \# \left\{ 0 \le j \le i - 1 : X(j) = \min_{j \le k \le i} X(k) \right\}.$

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Galton-Watson process

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- ► This individual has a number of children distributed according to the offspring distribution µ, where µ(k) gives the probability of k children, k ≥ 0.
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 Z_n gives the number of individuals in generation n (in particular, $Z_0 = 1$).

A Galton-Watson tree is the family tree arising from a Galton-Watson branching process.

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We will consider the case where the offspring distribution μ is critical or subcritical i.e.

$$\sum_{k=1}^{\infty} k\mu(k) \leq 1.$$

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This ensures that the resulting tree, T, is finite.

Since the tree is random, we will refer to the height *process* rather than function.

Uniform random trees revisited

Proposition

Let P be a (rooted, ordered) Galton-Watson tree, with Poisson(1) offspring distribution and total progeny N.

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Other combinatorial trees (in disguise)

Let T be a Galton-Watson tree with offspring distribution μ and total progeny N.

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Let T be a Galton-Watson tree with offspring distribution μ and total progeny N.

If µ(k) = 2^{-k-1}, k ≥ 0 (i.e. Geometric(1/2) offspring distribution) then conditional on N = n, the tree is uniform on the set of ordered trees with n vertices.

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- If µ(k) = ½(δ₀(k) + δ₂(k)), k ≥ 0 then conditional on N = n, for n odd, the tree is uniform on the set of (complete) binary trees.

Recall that μ is a distribution on \mathbb{Z}_+ such that $\sum_{k=1}^{\infty} k\mu(k) \leq 1$.

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Proposition

Let $(R(k), k \ge 0)$ be a random walk with initial value 0 and step distribution $\nu(k) = \mu(k+1), k \ge -1$. Set

$$M = \inf\{k \ge 0 : R(k) = -1\}.$$

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Now suppose that T is a Galton-Watson tree with offspring distribution μ and total progeny N. Then

$$(X(k), 0 \le k \le N) \stackrel{d}{=} (R(k), 0 \le k \le M).$$

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[Careful proof: see Le Gall.]

Galton-Watson forest

It turns out to be technically easier to deal with a sequence of i.i.d. Galton-Watson trees rather than a single tree. We can concatenate their height processes in order to encode the whole Galton-Watson forest.

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It can be checked that we still have

$$H(i) = \# \left\{ 0 \le j \le i - 1 : X(j) = \min_{j \le k \le i} X(k) \right\}, i \ge 0.$$

Convergence of the depth-first walk

Now specialise to the case where μ is critical and has finite offspring variance $\sigma^2 > 0$.

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Then $(X(k), k \ge 0)$ is a random walk with no drift and finite-variance step sizes.

Proposition (Donsker's theorem)

As
$$n \to \infty$$
, $\left(\frac{1}{\sqrt{n}}X(\lfloor nt \rfloor), t \ge 0\right) \stackrel{d}{\to} \sigma(B(t), t \ge 0),$

where $(B(t), t \ge 0)$ is a standard Brownian motion.

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Convergence of the height process

Theorem

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$$\left(rac{1}{\sqrt{n}} H(\lfloor nt
floor), t \geq 0
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Convergence of the height process: finite-dimensional distributions

Lemma

For any
$$m \ge 1$$
 and $0 \le t_1 \le t_2 \le \ldots \le t_m < \infty$,
 $\frac{1}{\sqrt{n}}(H(\lfloor nt_1 \rfloor), H(\lfloor nt_2 \rfloor), \ldots, H(\lfloor nt_m \rfloor)) \xrightarrow{d} \frac{2}{\sigma}(|B_{t_1}|, |B_{t_2}|, \ldots, |B_{t_m}|)$
as $n \to \infty$.

In order to get the functional convergence stated in the theorem, it remains to demonstrate that we have tightness. [Proof: see Le Gall.]

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Each excursion above 0 of the height process of the Galton-Watson forest corresponds to a tree, and the length of the excursion corresponds to the total progeny of that tree. If we condition the total progeny of the tree to be n, and let $n \to \infty$, intuitively we should obtain something like an excursion of the limit process.

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We need to make rigorous sense of what we mean by "an excursion of the limit process" before we can proceed.

A BRIEF INTRODUCTION TO EXCURSION THEORY

Partly based on **A guided tour through excursions** by Chris Rogers.

A tool: Itô's formula

Recall that for $f \in C^2(\mathbb{R}, \mathbb{R})$,

$$f(B_t) = f(B_0) + \int_0^t f'(B_s) dB_s + \frac{1}{2} \int_0^t f''(B_s) ds.$$

Consider a simple symmetric random walk $(X(n), n \ge 0)$. Let $T_0 = 0$ and, for $n \ge 1$,

$$T_n = \inf\{m > T_{n-1} : X(m) = 0\}.$$

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$$T_n = \inf\{m > T_{n-1} : X(m) = 0\}.$$

For $n \geq 1$, let

$$\xi^{n}(k) = \begin{cases} X(T_{n-1}+k) & \text{for } 0 \le k \le T_{n} - T_{n-1} \\ 0 & \text{for } k > T_{n} - T_{n-1}. \end{cases}$$

Then ξ^n is the *n*th excursion of X away from 0.

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By the Strong Markov property, ξ^1, ξ^2, \ldots are i.i.d.

In other words, the path of the random walk can be cut up into i.i.d. excursions away from 0.

Brownian excursions

Since the path of a Brownian motion $(B_t, t \ge 0)$ is continuous, the set $\{t : B_t \ne 0\}$ is open and so we can express it as a disjoint countable union of maximal open intervals $\bigcup_{i=1}^{\infty} (g_i, d_i)$ during which *B* makes an excursion away from 0.

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Let $\mathcal{Z} = \{t : B_t = 0\}$. It turns out to be essential to have a measure of how much time *B* spends at 0. The obvious one doesn't work:

Proposition

 $\operatorname{Leb}(\mathcal{Z}) = 0 \ a.s.$

The zero set of a Brownian motion

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- By the Strong Markov property, any point in Z is a limit of other points in Z.
- ▶ Z is closed by the continuity of the Brownian path.
- \mathcal{Z} is almost surely uncountable.

Think of \mathcal{Z} as being similar to the Cantor set (only random).

We want a process $(L_t)_{t\geq 0}$ which increases on \mathcal{Z} and is constant off it.

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We want a process $(L_t)_{t\geq 0}$ which increases on $\mathcal Z$ and is constant off it.

Definition (Brownian local time)

$$L_t = \lim_{\epsilon \downarrow 0} \frac{1}{2\epsilon} \int_0^t \mathbb{1}_{\{|B_s| \le \epsilon\}} ds.$$

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Note that $(L_t, t \ge 0)$ is clearly increasing.

Why is this the right definition?

Consider again the simple symmetric random walk on \mathbb{Z} , started

from 0. For
$$m \in \mathbb{Z}$$
, let $\operatorname{sgn}(m) = \begin{cases} 1 & \text{if } m > 0 \\ 0 & \text{if } m = 0 \\ -1 & \text{if } m < 0 \end{cases}$

Then for $n \geq 1$,

$$|X(n)| = \sum_{k=0}^{n-1} \operatorname{sgn}(X(k))(X(k+1) - X(k)) + \sum_{k=0}^{n-1} \mathbb{1}_{\{X(k)=0\}}$$

and so

$$\sum_{k=0}^{n-1} \mathbb{1}_{\{X(k)=0\}} = |X(n)| - \sum_{k=0}^{n-1} \operatorname{sgn}(X(k))(X(k+1) - X(k)).$$

This can be easily understood as an actual measure of how much time the random walk spends at the origin.

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This can be easily understood as an actual measure of how much time the random walk spends at the origin. Now imagine rescaling and using Donsker's theorem. There should be a limiting version of this equation for Brownian motion.

Tanaka's formula

Theorem

$$L_t = |B_t| - \int_0^t \operatorname{sgn}(B_s) dB_s.$$

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Note that this entails that $(L_t, t \ge 0)$ is continuous.

Local time measures the time spent at 0

Theorem

 $(L_t, t \ge 0)$ almost surely increases only on the set \mathcal{Z} .



Local time measures the time spent at 0

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 $(L_t, t \ge 0)$ almost surely increases only on the set \mathcal{Z} .

So $(L_t, t \ge 0)$ is constant during excursions away from 0.

Recall that we can write $\{t : B_t \neq 0\} = \bigcup_{i=1}^{\infty} (g_i, d_i)$. For each *i*, the excursion is $\xi^i = (B_{(g_i+t)\vee d_i}, t \ge 0)$, which takes values in

$$\mathcal{E} = \{ f \in \mathcal{C}([0,\infty),\mathbb{R}) : f(0) = 0, f(t) \neq 0 \text{ for } t \in (0,\zeta), \ f(t) = 0 \text{ for } t \geq \zeta, \text{ some } 0 < \zeta < \infty \}.$$

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The ordering cannot be captured by \mathbb{N} , but it turns out that it can be captured by the local time: we can think of the excursion straddling (g_i, d_i) as the excursion at local time ℓ for some ℓ , which occurs before the excursion straddling (g_j, d_j) , the excursion at local time $\ell' > \ell$.

Let $\tau_{\ell} = \inf\{t \ge 0 : L_t > \ell\}$. $(\tau_{\ell}, \ell \ge 0)$ is clearly right-continuous and increasing since $(L_t, t \ge 0)$ is continuous and increasing.

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There are only countably many values of ℓ such that $\Xi_{\ell} \neq \delta$, but there are infinitely many of them in (a, b) for $0 \leq a < b$.

A Poisson point process of excursions Theorem (Itô (1970))

 Ξ is a Poisson point process with intensity measure Leb \times n where n is a σ -finite measure on \mathcal{E} called the excursion measure.

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Proposition

$$\operatorname{m}({f \in \mathcal{E} : \zeta \ge x}) = \sqrt{\frac{2}{\pi x}}.$$

[See Kallenberg Foundations of modern probability for a nice proof.]

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It turns out that this carries over to its excursions.

Let
$$\mathcal{E}_x = \{f \in \mathcal{E} : \zeta = x\}$$
. For $f \in \mathcal{E}$ with duration ζ , put
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Proposition

For any $A \subseteq \mathcal{E}_{x}$,

$$\mathrm{n}(\nu_x^{-1}(A)|\zeta \geq c) := \frac{\mathrm{n}(\nu_x^{-1}(A) \cap \{\zeta \geq c\})}{\mathrm{n}(\zeta \geq c)}$$

does not depend on c > 0.

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A little more work shows that we can make sense of $n^{(x)}(A) := n(|f| \in A | \zeta = x)$ as a probability measure on $\mathcal{E}_x^+ = \{f \in \mathcal{E}_x : f \ge 0\}$, the law of a process called a Brownian excursion of length x,

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We refer to $(e^{(1)}(t), 0 \le t \le 1)$ as a standard Brownian excursion (and usually omit the superscript in this case).

Standard Brownian excursion, $(e(t), 0 \le t \le 1)$



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Two-stage description of the excursion process

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► Take a Poisson point process Θ on $[0, \infty) \times [0, \infty)$ of intensity Leb × m, where $m(dx) = n(\zeta \in dx) = (2\pi)^{-1/2} x^{-3/2} dx$.

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- ► Take a Poisson point process Θ on $[0, \infty) \times [0, \infty)$ of intensity Leb × m, where $m(dx) = n(\zeta \in dx) = (2\pi)^{-1/2} x^{-3/2} dx$.
- For a point at (ℓ, ζ), sample a standard Brownian excursion e_ℓ. Then (√ζe_ℓ(t/ζ), t ≥ 0) gives the excursion straddling local time ℓ.

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Some loose ends: Galton-Watson trees

Recall that we showed that a Galton-Watson forest can be coded by its depth-first walk and height process.

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Recall that we showed that a Galton-Watson forest can be coded by its depth-first walk and height process. We showed that as $n \to \infty$,

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Galton-Watson trees conditioned on their total progeny

Recall that the depth-first walk X of a critical Galton-Watson tree with offspring variance $\sigma^2 > 0$ is a random walk with step mean 0 and variance σ^2 . The total progeny N is equal to $\inf\{k \ge 0 : X(k) = -1\}$. Write $(X^n(k), 0 \le k \le n)$ for the depth-first walk conditioned on N = n.

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Lemma

As $n \to \infty$,

$$(n^{-1/2}X^n(\lfloor nt \rfloor), 0 \le t \le 1) \xrightarrow{d} \sigma(e(t), 0 \le t \le 1).$$

[See W.D. Kaigh **An invariance principle for random walk conditioned by a late return to zero** *Annals of Probability* **4** (1976) pp.115-121.]

Convergence of the coding processes

Let $(X^n(i), 0 \le i \le n)$ and $(H^n(i), 0 \le i \le n)$ be the depth-first walk and height process respectively of a critical Galton-Watson tree with offspring variance $\sigma^2 > 0$, conditioned to have total progeny *n*.

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As $n \to \infty$, $(n^{-1/2}X^n(\lfloor n \cdot \rfloor), n^{-1/2}H^n(\lfloor n \cdot \rfloor)) \xrightarrow{d} \left(\sigma e, \frac{2}{\sigma}e\right)$, where $e = (e(t), 0 \le t \le 1)$ is a standard Brownian excursion.

[Proof: see Le Gall.]

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where $e = (e(t), 0 \le t \le 1)$ is a standard Brownian excursion.

[Proof: see Le Gall.] This result suggests the existence of some sort of limiting tree, which is "coded" by the Brownian excursion.

REAL TREES

Real trees

Definition

A compact metric space (\mathcal{T}, d) is a real tree if for all $x, y \in \mathcal{T}$,

► There exists a unique shortest path [[x, y]] from x to y (of length d(x, y)).

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An element $v \in T$ is called a vertex. A rooted real tree has a distinguished vertex ρ called the root. The height of a vertex v is its distance $d(\rho, v)$ from the root. A leaf is a vertex v such that $v \notin [[\rho, w]]$ for any $w \neq v$.

Suppose that $h: [0, \infty) \to [0, \infty)$ is a continuous function of compact support such that h(0) = 0. *h* will play the role of the height process for a real tree.

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Use h to define a distance:

$$d_h(x,y) = h(x) + h(y) - 2 \inf_{x \land y \le z \le x \lor y} h(z).$$

Let $y \sim y'$ if $d_h(y,y') = 0$ and take the quotient $\mathcal{T}_h = [0,\infty)/\sim$.



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Theorem

 (\mathcal{T}_h, d_h) is a real tree.

[Proof: see Le Gall.]

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We will always take the equivalence class of 0 to be the root, ρ .

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Theorem

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We will always take the equivalence class of 0 to be the root, ρ .

Definition

The Brownian continuum random tree is T_{2e} , where e is a standard Brownian excursion.

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The Brownian continuum random tree T_{2e}



[Picture by Grégory Miermont]

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The Hausdorff distance between two compact subsets K and K' of a metric space (M, δ) is

$$d_H(K,K') = \inf\{\epsilon > 0 : K \subseteq F_\epsilon(K'), K' \subseteq F_\epsilon(K)\},$$

where $F_{\epsilon}(K) := \{x \in M : \delta(x, K) \leq \epsilon\}$ is the ϵ -fattening of K.



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To measure the distance between two compact metric spaces (X, d) and (X', d'), the idea is to embed them (isometrically) into a single larger metric space and then compare them using the Hausdorff distance.

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To measure the distance between two compact metric spaces (X, d) and (X', d'), the idea is to embed them (isometrically) into a single larger metric space and then compare them using the Hausdorff distance.

So define the Gromov-Hausdorff distance

$$d_{GH}(X,X') = \inf\{d_H(\phi(X),\phi'(X'))\},\$$

where the infimum is taken over all choices of metric space (M, δ) and all isometric embeddings $\phi : X \to M$, $\phi' : X' \to M$.

If the metric spaces are rooted, at ρ and ρ' respectively, we take

$$d_{GH}(X,X') = \inf\{d_H(\phi(X),\phi'(X')) \lor \delta(\phi(\rho),\phi'(\rho'))\}$$

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Fortunately, we do not have to seek an optimal embedding!

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For compact metric spaces (X, d) and (X', d'), a correspondence between X and X' is a subset \mathcal{R} of $X \times X'$ such that for each $x \in X$, there exists at least one $x' \in X'$ such that $(x, x') \in \mathcal{R}$ and vice versa.

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The distortion of a correspondence $\ensuremath{\mathcal{R}}$ is defined by

$$\operatorname{dis}(\mathcal{R}) = \sup\{|d(x,y) - d'(x',y')| : (x,x'), (y,y') \in \mathcal{R}\}.$$

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Proposition

If X and X' are compact metric spaces rooted at ρ and ρ' respectively then

$$d_{GH}(X,X') = \frac{1}{2}\inf\operatorname{dis}(\mathcal{R}),$$

where the infimum is taken over all correspondences \mathcal{R} between X and X' such that $(\rho, \rho') \in \mathcal{R}$.

Convergence to the CRT

Let T_n be our Galton-Watson tree conditioned to have size n.

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Theorem

As
$$n o \infty$$
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where convergence is in the Gromov-Hausdorff sense.

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[Approach due to Grégory Miermont.]

The mass measure of the CRT

Consider now a uniform random tree T_n . Put mass 1/n at each vertex. Call the resulting probability measure μ_n . It should be intuitively clear that

$$\left(\frac{1}{\sqrt{n}}T_n,\mu_n\right) \stackrel{d}{\rightarrow} (\mathcal{T}_{2e},\mu),$$

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Lemma

Let \mathcal{L} be the set of leaves of \mathcal{T}_{2e} . Then

$$\mu(\mathcal{L}) = 1.$$

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Lemma

Let \mathcal{L} be the set of leaves of \mathcal{T}_{2e} . Then

$$\mu(\mathcal{L}) = 1.$$

[Intuition: non-leaf vertices of T_n are typically at distance $o(\sqrt{n})$ from a leaf. Proof: see Aldous (1991).]
RANDOM GRAPHS

The Erdős-Rényi random graph

Take *n* vertices labelled by $[n] := \{1, 2, ..., n\}$ and put an edge between any pair independently with probability *p*. Call the resulting model G(n, p).

Example: n = 10, p = 0.4 (vertex labels omitted).



Connected components

We're going to be interested in the connected components of these graphs.

Below, there are three of them.



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The phase transition

Let p = c/n and consider the largest component (vertices in green, edges in red).

n = 200, c = 0.4



The phase transition

Let p = c/n and consider the largest component (vertices in green, edges in red).

n = 200, c = 0.8



The phase transition

Let p = c/n and consider the largest component (vertices in green, edges in red).

n = 200, c = 1.2



The phase transition (Erdős and Rényi (1960))

By the size of a component, we mean its number of vertices.

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The phase transition (Erdős and Rényi (1960))

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Consider p = c/n.

For c < 1, the largest connected component has size O(log n);
for c > 1, the largest connected component has size Θ(n) (and the others are all O(log n)).

[These statements hold with probability tending to 1 as $n \to \infty$.]

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[These statements hold with probability tending to 1 as $n \to \infty$.]

If c = 1, the largest component has size $\Theta(n^{2/3})$ and, indeed, there is a whole sequence of components of this order.

The critical random graph

The critical window: $p = \frac{1}{n} + \frac{\lambda}{n^{4/3}}$, where $\lambda \in \mathbb{R}$. For such p, the largest components have size $\Theta(n^{2/3})$.

The critical random graph

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We will also be interested in the surplus of a component, the number of edges more than a tree that it has.

A component with surplus 3:



Convergence of the sizes and surpluses

Fix λ and let C_1^n, C_2^n, \ldots be the sequence of component sizes in decreasing order, and let S_1^n, S_2^n, \ldots be their surpluses.

Write $\mathbf{C}^n = (C_1^n, C_2^n, ...)$ and $\mathbf{S}^n = (S_1^n, S_2^n, ...)$.

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Write $C^n = (C_1^n, C_2^n, ...)$ and $S^n = (S_1^n, S_2^n, ...)$.

Theorem (Aldous (1997))

As
$$n \to \infty$$
, $(n^{-2/3}\mathbf{C}^n, \mathbf{S}^n) \stackrel{d}{\to} (\mathbf{C}, \mathbf{S}).$

Let $W^{\lambda}(t) = W(t) + \lambda t - \frac{t^2}{2}$, $t \ge 0$, where $(W(t), t \ge 0)$ is a standard Brownian motion.

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Let $B^{\lambda}(t) = W^{\lambda}(t) - \min_{0 \le s \le t} W^{\lambda}(s)$ be the process reflected at its minimum.



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Decorate the picture with the points of a rate one Poisson process which fall above the *x*-axis and below the graph.

 ${\bf C}$ is the sequence of excursion-lengths of this process, in decreasing order.

 ${\boldsymbol{\mathsf{S}}}$ is the sequence of numbers of points falling in the corresponding excursions.

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Convergence of the sizes and surpluses

Theorem (Aldous (1997))

As
$$n \to \infty$$
, $(n^{-2/3} \mathbf{C}^n, \mathbf{S}^n) \stackrel{d}{\to} (\mathbf{C}, \mathbf{S}),$

where **C** is the sequence of excursion-lengths of B^{λ} in decreasing order, and **S** is the sequence of numbers of Poisson points falling in the corresponding excursions.

Convergence of the sizes and surpluses

Theorem (Aldous (1997))

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where **C** is the sequence of excursion-lengths of B^{λ} in decreasing order, and **S** is the sequence of numbers of Poisson points falling in the corresponding excursions.

Here, convergence in the first co-ordinate takes place in

$$\ell^2_{\searrow} := \left\{ \mathbf{x} = (x_1, x_2, \ldots) : x_1 \ge x_2 \ge \ldots \ge 0, \sum_{i=1}^{\infty} x_i^2 < \infty \right\}.$$

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Proof technique: depth-first exploration

As for the discrete trees at the beginning of the course, a key tool is a depth-first exploration. We previously defined the depth-first walk by X(0) = 0 and, for $1 \le k \le n$,

$$X(k) = \sum_{i=0}^{k-1} (k(v_i) - 1),$$

where k(v) is the number of children of vertex v and $v_0, v_1, \ldots, v_{n-1}$ are the vertices in lexicographical order.

There are two problems with this definition: the components of a random graph are are labelled but not ordered, and they are not (in general) trees.

These problems are resolved by stepping through the graph vertex by vertex, using the natural ordering of the labels, and ignoring non-tree edges. Exactly how we do this is best explained on an example.

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It's useful to say that vertices can have four states: current, alive, dead or unexplored. For the first component, X(k) will turn out to be the number of alive vertices at step k. Thereafter, it will be the number of vertices alive minus the number of components already fully explored.

Step 0



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Current: 1 Alive: none Dead: none X(0) = 0.

Step 1



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Current: 5 Alive: 7, 10 Dead: 1 X(1) = 2.

Step 2



Current: 2 Alive: 9, 7, 10 Dead: 1, 5 X(2) = 3.

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Step 3



Current: 3 Alive: 9, 7, 10 Dead: 1, 5, 2 X(3) = 3.

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Step 4



Current: 9 Alive: 7, 10 Dead: 1, 5, 2, 3 X(4) = 2.

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Step 5



Current: 7 Alive: 10 Dead: 1, 5, 2, 3, 9 X(5) = 1.

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Step 6



Current: 10 Alive: none Dead: 1, 5, 2, 3, 9, 7 X(6) = 0.

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Step 7



Current: 8 Alive: none Dead: 1, 5, 2, 3, 9, 7, 10 X(7) = 0.

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Step 8



Current: 4 Alive: 6 Dead: 1, 5, 2, 3, 9, 7, 10, 8 X(8) = 1.

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Step 9



Current: 6 Alive: none Dead: 1, 5, 2, 3, 9, 7, 10, 8, 4 X(9) = 0.

We explored the graph on the left as if it were the tree on the right:



Depth-first walk



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Depth-first walk



As for a forest, if there are several components, $T(k) = \inf\{i \ge 0 : X(i) = -k\}$ marks the beginning of the (k+1)th component. So the component sizes are $\{T(k+1) - T(k), k \ge 0\}$. This sequence can clearly be reconstructed from the path of $(X(i), i \ge 0)$.
Convergence of the depth-first walk

Let X_n^{λ} be the depth-first walk associated with $G(n, n^{-1} + \lambda n^{-4/3})$.

Theorem

As $n \to \infty$, $(n^{-1/3}X_n^{\lambda}(\lfloor n^{2/3}t \rfloor), t \ge 0) \xrightarrow{d} (W^{\lambda}(t), t \ge 0).$

The convergence here is uniform on compact time-intervals.

To finish the proof...

A little care needs to be taken to check that the lengths of excursions above past-minima of X_n^{λ} converge to lengths of excursions above past-minima of W^{λ} , and that we don't miss any excursions of length $\Omega(n^{2/3})$. [Proof: see Aldous (1997).]

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We will deal with the surplus edges a little later.



What do the limiting components look like?



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The vertex-labels are irrelevant: we are really interested in what distances look like in the limit. So we will give a metric space answer, and convergence will be in the Gromov-Hausdorff distance.

Simple but important fact: a component of G(n, p) conditioned to have *m* vertices and *s* surplus edges is a uniform connected graph on those *m* vertices with m + s - 1 edges.

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Our general approach is to pick out a spanning tree, and then to put in the surplus edges.

Depth-first tree

In the depth-first exploration, we effectively explored this spanning tree; the dashed edges made no difference.



Call it the depth-first tree associated with the graph G, and write T(G).

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There is one case which we already understand: when the surplus of a component is 0. Then the component is a uniform random tree (and is necessarily the same as its depth-first tree). In this case, it is clear that the scaling limit is the Brownian CRT.

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In the tree case, we should rescale distances by $1/\sqrt{m}$, where *m* is the number of vertices in the component. This is the correct distance rescaling for all of the big components in the random graph.

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Each excursion of the process $(B^{\lambda}(t), t \ge 0)$ of length x corresponds to the limit of a component on $\sim xn^{2/3}$ vertices. Such an excursion codes a continuum random tree, which is a "spanning tree" for that limit component.

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In the limit, surplus edges correspond to vertex-identifications (since edge-lengths have shrunk to 0). In each excursion, the points of the Poisson process tell us where these vertex-identifications should occur.

Excursions of the limit process

Consider the process $(B^{\lambda}(t), t \ge 0)$.

Lemma

An excursion $\tilde{e}^{(x)}$ of $(B^{\lambda}(t), t \ge 0)$, conditioned to have length x, has a distribution specified by

$$\mathbb{E}\left[f\left(\tilde{e}^{(x)}\right)\right] = \frac{\mathbb{E}\left[f\left(e^{(x)}\right)\exp\left(\int_{0}^{x}e^{(x)}(u)du\right)\right]}{\mathbb{E}\left[\exp\left(\int_{0}^{x}e^{(x)}(u)du\right)\right]},$$

where f is any suitable test-function and $e^{(x)}$ is a Brownian excursion of length x.

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where f is any suitable test-function and $e^{(x)}$ is a Brownian excursion of length x.

Note that this holds independently of λ . We refer to $\tilde{e}^{(x)}$ as a tilted excursion and to the tree \tilde{T} that it encodes as a tilted tree.

Vertex identifications



A point at (x, y) identifies the vertex v at height h(x) with the vertex at distance y along the path from the root to v.

Note that it follows from properties of the tilted trees and of the Poisson process that we may equivalently describe the limit of a component on $\sim xn^{2/3}$ vertices as follows.

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Sample a tilted excursion $\tilde{\mathbf{e}}^{(x)}$ of length x and use it to create a CRT $\tilde{\mathcal{T}}.$



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Conditional on $\tilde{e}^{(x)}$, sample a random variable P with Poisson $\left(\int_{0}^{x} \tilde{e}^{(x)}(u) du\right)$ distribution.

Conditional on P = s, pick s vertices of the tree \tilde{T} independently with density proportional to their height. (These will almost surely be leaves.)



For each of the selected leaves, pick a uniform point on the path from the leaf to the root.



Identify each of the selected leaves with its chosen point.



Convergence result

Let C_1^n, C_2^n, \ldots be the sequence of components of G(n, p) in decreasing order of size, considered as metric spaces with the graph distance.

Theorem

As
$$n \to \infty$$
,
 $n^{-1/3}(\mathcal{C}_1^n, \mathcal{C}_2^n, \ldots) \xrightarrow{d} (\mathcal{C}_1, \mathcal{C}_2, \ldots)$

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where C_1, C_2, \ldots is the sequence of metric spaces corresponding to the excursions of the marked limit process B^{λ} in decreasing order of length.

Here, convergence is with respect to the metric

$$d(\mathcal{A},\mathcal{B}) := \left(\sum_{i=1}^{\infty} d_{GH}(\mathcal{A}_i,\mathcal{B}_i)^4\right)^{1/4}.$$

Idea of proof

The key idea turns out to be study a component of G(n, p) conditioned on its size but *not* on its surplus.

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Depth-first tree

Take an arbitrary component G of G(n, p). Recall that T(G) is the depth-first tree associated with G



and that $(X(k), 0 \le k \le n)$ is the depth-first walk of T(G).

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Permitted edges

Look at things the other way round: for a given tree T, which connected graphs G have depth-first tree T(G) = T? In other words, where can we put surplus edges so that they don't change T?

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Call such edges permitted.



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Step 0: X(0) = 0.



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Step 1: X(1) = 2.



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Step 2: X(2) = 3.



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Step 3: X(3) = 3.



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Step 4: X(4) = 2.



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Step 5: X(5) = 1.



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Step 6: X(6) = 0.



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Step 7: X(7) = 0.
Depth-first walk and permitted edges



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Step 8: X(8) = 1.

Depth-first walk and permitted edges



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Step 10: X(9) = 0.

Area

At step $k \ge 0$ there are X(k) permitted edges. So the total number is

$$a(T) = \sum_{k=0}^{m-1} X(k).$$

We call this the area of T.



Classifying graphs by depth-first tree

Let \mathbb{G}_T be the set of graphs G such that T(G) = T. It follows that $|\mathbb{G}_T| = 2^{a(T)}$, since each permitted edge may either be included or not.

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Classifying graphs by depth-first tree

Let \mathbb{G}_T be the set of graphs G such that T(G) = T. It follows that $|\mathbb{G}_T| = 2^{a(T)}$, since each permitted edge may either be included or not.

Recall that $\mathbb{T}_{[m]}$ is the set of trees with label-set $[m] = \{1, 2, \dots, m\}$. Then

$$\left\{\mathbb{G}_T: T \in \mathbb{T}_{[m]}\right\}$$

is a partition of the set of connected graphs on [m].

Create a connected graph \tilde{G}_m^p as follows.

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• Pick a random labelled tree \tilde{T}_m^p such that

$$\mathbb{P}\left(\tilde{T}_m^p=T
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Create a connected graph \tilde{G}_m^p as follows.

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• Add each of the $a(\tilde{T}_m^p)$ permitted edges to \tilde{T}_m^p independently with probability p.

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► Add each of the a(T^p_m) permitted edges to T^p_m independently with probability p.

Lemma

 \tilde{G}_m^p has the same distribution as G_m^p , a component of G(n, p) conditioned to have vertex-set [m].

Taking limits

So we need to prove that

- ▶ the tree \tilde{T}_m^p converges to a CRT coded by a tilted excursion;
- the locations of the surplus edges converge to the locations in our limiting picture.

We will deal with the tree first. For simplicity, we will take $p = m^{-3/2}$; the general case is similar.

Convergence of the tree

Theorem

Suppose $p = m^{-3/2}$. Then

$$\frac{1}{\sqrt{m}}\tilde{T}^p_m \stackrel{d}{\to} \tilde{T}$$

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as $m \to \infty$.

The permitted edges are in bijective correspondence with the integer points under the graph of the depth-first walk.



The permitted edges are in bijective correspondence with the integer points under the graph of the depth-first walk. Since each permitted edge is included independently with probability p, the surplus edges form a Binomial point process.



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A point at (k, j) means "put an edge between the current vertex at step k and the vertex at distance j from the bottom of the list of alive vertices".

Surplus edges almost go to ancestors... In fact, they always go to younger children of ancestors of the current vertex.



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When we rescale, the distance between a vertex and one of its children vanishes and so, in the limit, surplus "edges" do go to ancestors of the current vertex.

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The Binomial point process of surplus edges, when rescaled, straightforwardly converges to the required Poisson point process.

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