

Supervised Learning

Unsupervised learning:

- ▶ To “extract structure” and postulate hypotheses about data generating process from observations x_1, \dots, x_n .
- ▶ Visualize, summarize and compress data.

We have seen how response or grouping variables are used to validate the usefulness of the extracted structure.

Supervised learning:

- ▶ In addition to the n observations of X , we also have a response variable $Y \in \mathcal{Y}$.
- ▶ Techniques for predicting Y given X .
 - ▶ Classification: discrete responses, e.g. $\mathcal{Y} = \{+1, -1\}$ or $\{1, \dots, K\}$.
 - ▶ Regression: a numerical value is observed and $\mathcal{Y} = \mathbb{R}$.

Given training data (x_i, y_i) , $i = 1, \dots, n$, the goal is to accurately predict the class or response Y on new observations of X .

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Regression Example: Boston Housing

The original data are 506 observations on 13 variables X ; medv being the response variable Y .

```
crim    per capita crime rate by town
zn      proportion of residential land zoned for lots
        over 25,000 sq.ft
indus   proportion of non-retail business acres per town
chas    Charles River dummy variable (= 1 if tract bounds river;
        0 otherwise)
nox     nitric oxides concentration (parts per 10 million)
rm      average number of rooms per dwelling
age     proportion of owner-occupied units built prior to 1940
dis     weighted distances to five Boston employment centers
rad     index of accessibility to radial highways
tax     full-value property-tax rate per USD 10,000
ptratio pupil-teacher ratio by town
b       1000(B - 0.63)^2 where B is the proportion of blacks by town
lstat   percentage of lower status of the population
medv    median value of owner-occupied homes in USD 1000's
```

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Regression Example: Boston Housing

```
> str(X)
'data.frame':  506 obs. of  13 variables:
 $ crim   : num  0.00632 0.02731 0.02729 0.03237 0.06905 ...
 $ zn     : num  18 0 0 0 0 0 12.5 12.5 12.5 ...
 $ indus  : num  2.31 7.07 7.07 2.18 2.18 2.18 7.87 7.87 7.87 7.87
 $ chas   : int  0 0 0 0 0 0 0 0 0 ...
 $ nox    : num  0.538 0.469 0.469 0.458 0.458 0.458 0.524 0.524 ...
 $ rm     : num  6.58 6.42 7.18 7.00 7.15 ...
 $ age    : num  65.2 78.9 61.1 45.8 54.2 58.7 66.6 96.1 100 85.9
 $ dis    : num  4.09 4.97 4.97 6.06 6.06 ...
 $ rad    : int  1 2 2 3 3 3 5 5 5 5 ...
 $ tax    : num  296 242 242 222 222 222 311 311 311 311 ...
 $ ptratio: num  15.3 17.8 17.8 18.7 18.7 18.7 15.2 15.2 15.2 15.2
 $ black  : num  397 397 393 395 397 ...
 $ lstat  : num  4.98 9.14 4.03 2.94 5.33 ...
```

```
> str(Y)
num[1:506] 24 21.6 34.7 33.4 36.2 28.7 22.9 27.1 16.5 18.9 ...
```

Goal: predict median house price $\hat{Y}(X)$, given 13 predictor variables X of a new district.

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Classification Example: Lymphoma

We have gene expression measurements X of $n = 62$ patients for $p = 4026$ genes. For each patient, Y denotes one of two subtypes of cancer. Goal: predict cancer subtype $\hat{Y}(X) \in \{0, 1\}$, given gene expressions of a new patient.

```
> str(X)
'data.frame':  62 obs. of  4026 variables:
 $ Gene 1 : num -0.344 -1.188 0.520 -0.748 -0.868 ...
 $ Gene 2 : num -0.953 -1.286 0.657 -1.328 -1.330 ...
 $ Gene 3 : num -0.776 -0.588 0.409 -0.991 -1.517 ...
 $ Gene 4 : num -0.474 -1.588 0.219 0.978 -1.604 ...
 $ Gene 5 : num -1.896 -1.960 -1.695 -0.348 -0.595 ...
 $ Gene 6 : num -2.075 -2.117 0.121 -0.800 0.651 ...
 $ Gene 7 : num -1.8755 -1.8187 0.3175 0.3873 0.0414 ...
 $ Gene 8 : num -1.539 -2.433 -0.337 -0.522 -0.668 ...
 $ Gene 9 : num -0.604 -0.710 -1.269 -0.832 0.458 ...
 $ Gene 10: num -0.218 -0.487 -1.203 -0.919 -0.848 ...
 $ Gene 11: num -0.340 1.164 1.023 1.133 -0.541 ...
 $ Gene 12: num -0.531 0.488 -0.335 0.496 -0.358 ...
```

```
> str(Y)
num [1:62] 0 0 0 1 0 0 1 0 0 0 ...
```

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Decision Theory

- ▶ Suppose we made a prediction $\hat{Y} \in \mathcal{Y}$ based on observation of X .
- ▶ How good is the prediction? We can use a **loss function** $L : \mathcal{Y} \times \mathcal{Y} \mapsto \mathbb{R}^+$ to formalize the quality of the prediction.
- ▶ Typical loss functions:
 - ▶ **Misclassification loss** (or **0-1 loss**) for classification

$$L(Y, \hat{Y}) = \begin{cases} 0 & Y = \hat{Y} \\ 1 & Y \neq \hat{Y} \end{cases}.$$

- ▶ **Squared loss** for regression

$$L(Y, \hat{Y}) = (Y - \hat{Y})^2.$$

- ▶ Alternative loss functions are often useful (later). For example, **weighted misclassification error** often appropriate. Or **log-likelihood loss** (sometimes shortened as **log loss**) $L(Y, \hat{p}) = -\log \hat{p}(Y)$, where $\hat{p}(k)$ is the estimated probability of class $k \in \mathcal{Y}$.

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Decision Theory

- ▶ For a given loss function L , the **risk** R of a learner is given by the expected loss

$$R(\hat{Y}) = \mathbb{E}(L(Y, \hat{Y}(X))),$$

where the expectation is with respect to the true (unknown) joint distribution (X, Y) .

- ▶ The risk is unknown, but we can estimate it by the **empirical risk**:

$$R(\hat{Y}) \approx R_n(\hat{Y}) = \frac{1}{n} \sum_{i=1}^n L(y_i, \hat{Y}(x_i)).$$

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The Bayes Classifier

- ▶ What is the optimal classifier if the joint distribution (X, Y) were known?
- ▶ The joint distribution f of X can be written as a mixture

$$f(X) = \sum_{k=1}^K f_k(X) \mathbb{P}(Y = k),$$

where, for $k = 1, \dots, K$,

- ▶ the prior probabilities over classes are $P(Y = k) = \pi_k$
- ▶ and distributions of X , conditional on $Y = k$, is $f_k(X)$.

- ▶ The **Bayes classifier** $\hat{Y}(X) \mapsto \{1, \dots, K\}$ is the one with minimum risk:

$$\begin{aligned} R(\hat{Y}) &= \mathbb{E}[L(Y, \hat{Y}(X))] = \mathbb{E}[\mathbb{E}[L(Y, \hat{Y}(x)) | X = x]] \\ &= \int_{\mathcal{X}} \mathbb{E}[L(Y, \hat{Y}(x)) | X = x] f(x) dx \end{aligned}$$

- ▶ The minimum risk attained by the Bayes classifier is called **Bayes risk**.
- ▶ Minimizing $\mathbb{E}[L(Y, \hat{Y}(x)) | X = x]$ separately for each x suffices.

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The Bayes Classifier

- ▶ Consider the situation of the 0-1 loss.
- ▶ The risk simplifies to:

$$\begin{aligned} \mathbb{E}[L(Y, \hat{Y}(x)) | X = x] &= \sum_{k=1}^K L(k, \hat{Y}(x)) \mathbb{P}(Y = k | X = x) \\ &= 1 - \mathbb{P}(Y = \hat{Y}(x) | X = x) \end{aligned}$$

- ▶ The risk is minimized by choosing the class with the greatest posterior probability:

$$\begin{aligned} \hat{Y}(x) &= \arg \max_{k=1, \dots, K} \mathbb{P}(Y = k | X = x) = \arg \max_{k=1, \dots, K} \frac{\pi_k f_k(x)}{\sum_{k=1}^K \pi_k f_k(x)} \\ &= \arg \max_{k=1, \dots, K} \pi_k f_k(x). \end{aligned}$$

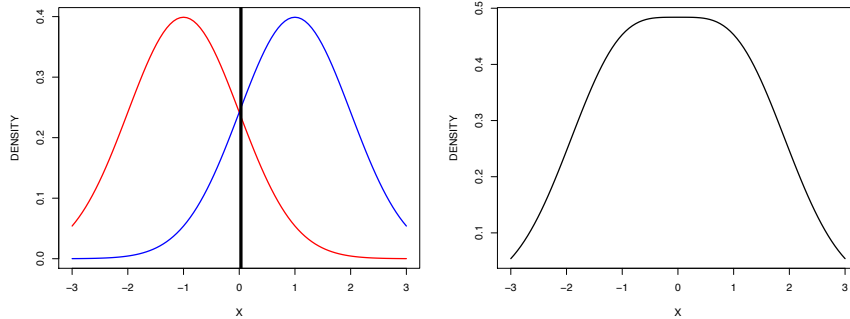
- ▶ The functions $x \mapsto \pi_k f_k(x)$ are called **discriminant functions**. The function with maximum value determines the predicted class of x .

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The Bayes Classifier

A simple two Gaussians example: Suppose $X \sim \mathcal{N}(\mu_Y, 1)$, where $\mu_1 = -1$ and $\mu_2 = 1$ and assume equal priors $\pi_1 = \pi_2 = 1/2$.

$$f_1(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(x - (-1))^2}{2}\right) \quad \text{and} \quad f_2(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(x - 1)^2}{2}\right).$$

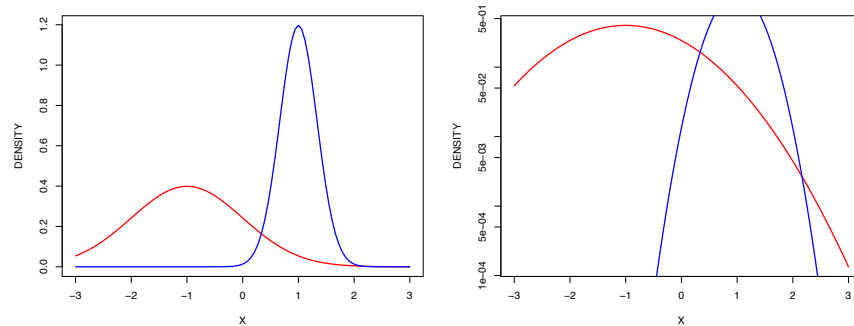


Optimal classification is $\hat{Y}(x) = \arg \max_{k=1, \dots, K} \pi_k f_k(x) = \begin{cases} 1 & \text{if } x < 0, \\ 2 & \text{if } x \geq 0. \end{cases}$

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The Bayes Classifier

How do you classify a new observation x if now the standard deviation is still 1 for class 1 but 1/3 for class 2?



Looking at density in a log-scale, optimal classification is class 2 if and only if $x \in [-0.39, 2.15]$.

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Plug-in Classification

- ▶ The Bayes Classifier chooses the class with the greatest posterior probability

$$\hat{Y}(x) = \arg \max_{k=1, \dots, K} \pi_k f_k(x).$$

- ▶ Unfortunately, we usually know neither the conditional class probabilities nor the prior probabilities.
- ▶ We can estimate the joint distribution with:
 - ▶ estimates $\hat{\pi}_k$ for π_k and $k = 1, \dots, K$ and
 - ▶ estimates $\hat{f}_k(x)$ of conditional class densities,
- ▶ The **plug-in classifiers** chooses the class

$$\hat{Y}(x) = \arg \max_{k=1, \dots, K} \hat{\pi}_k \hat{f}_k(x).$$

- ▶ **Linear Discriminant Analysis** will be an example of plug-in classification.

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Linear Discriminant Analysis

- ▶ LDA is the most well-known and simplest example of plug-in classification.
- ▶ Assume a multivariate Normal form for $f_k(x)$ for each class k :

$$X|Y = k \sim \mathcal{N}(\mu_k, \Sigma),$$

- ▶ each class can have a **different mean** μ_k
 - ▶ but all classes share the **same covariance** Σ .
- ▶ For an observation x ,

$$\begin{aligned} \log \mathbb{P}(Y = k|X = x) &= \kappa + \log \pi_k f_k(x) \\ &= \kappa + \log \pi_k - \frac{1}{2}(x - \mu_k)^\top \Sigma^{-1}(x - \mu_k) \end{aligned}$$

The quantity $(x - \mu_k)^\top \Sigma^{-1}(x - \mu_k)$ is the square of the **Mahalanobis distance**. It gives the distance between x and μ_k in the metric given by Σ .

- ▶ If $\Sigma = I_p$ and $\pi_k = \frac{1}{K}$, $\hat{Y}(x)$ simply chooses the class k with the nearest (in the Euclidean sense) mean.

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Linear Discriminant Analysis

- ▶ Expanding the **discriminant** $(x - \mu_k)^\top \Sigma^{-1} (x - \mu_k)$,

$$\begin{aligned} \log \mathbb{P}(Y = k|x) &= \kappa + \log(\pi_k) - \frac{1}{2} (\mu_k^\top \Sigma^{-1} \mu_k - 2\mu_k^\top \Sigma^{-1} x + x^\top \Sigma^{-1} x) \\ &= \kappa + \log(\pi_k) - \frac{1}{2} \mu_k^\top \Sigma^{-1} \mu_k + \mu_k^\top \Sigma^{-1} x \end{aligned}$$

- ▶ Setting $a_k = \log(\pi_k) - \frac{1}{2} \mu_k^\top \Sigma^{-1} \mu_k$ and $b_k = \Sigma^{-1} \mu_k$, we obtain

$$\log \mathbb{P}(Y = k|X = x) = \kappa + a_k + b_k^\top x$$

i.e. a **linear** discriminant function.

- ▶ Consider choosing class k over k' :

$$a_k + b_k^\top x > a_{k'} + b_{k'}^\top x \quad \Leftrightarrow \quad a_\star + b_\star^\top x > 0$$

where $a_\star = a_k - a_{k'}$ and $b_\star = b_k - b_{k'}$.

- ▶ The Bayes classifier partitions \mathcal{X} into regions with the same class predictions via **separating hyperplanes**.
- ▶ The Bayes classifier under these assumptions is more commonly known as the **LDA classifier**.

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Parameter Estimation

- ▶ The final piece of the puzzle is to estimate the parameters of the LDA model.
- ▶ We can achieve this by maximum likelihood.
- ▶ EM algorithm is not needed here since the class variables y_i are observed.
- ▶ Let $n_k = \#\{j : y_j = k\}$ be the number of observations in class k .

$$\ell(\pi, (\mu_k), \Sigma) = \kappa + \sum_{k=1}^K \sum_{j:y_j=k} \log \pi_k - \frac{1}{2} (\log |\Sigma| + (x_j - \mu_k)^\top \Sigma^{-1} (x_j - \mu_k))$$

Then:

$$\hat{\pi}_k = \frac{n_k}{n} \quad \hat{\mu}_k = \frac{1}{n_k} \sum_{j:y_j=k} x_j$$

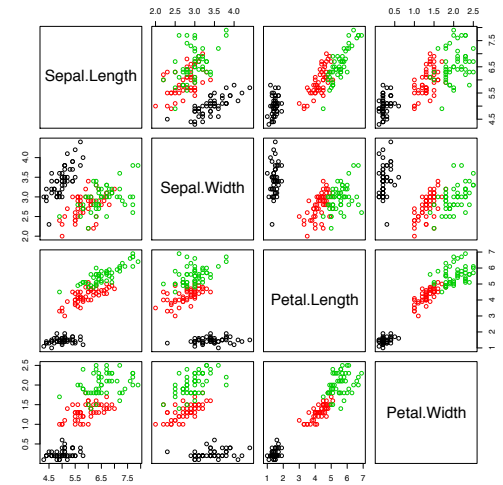
$$\hat{\Sigma} = \frac{1}{n} \sum_{k=1}^K \sum_{j:y_j=k} (x_j - \hat{\mu}_k)(x_j - \hat{\mu}_k)^\top$$

- ▶ Note: the ML estimate of Σ is not unbiased. For an unbiased estimate we need to divide by $n - K$.

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Iris Dataset

```
library(MASS)
data(iris)
##save class labels
ct <- rep(1:3,each=50)
##pairwise plot
pairs(iris[,1:4],col=ct)
```

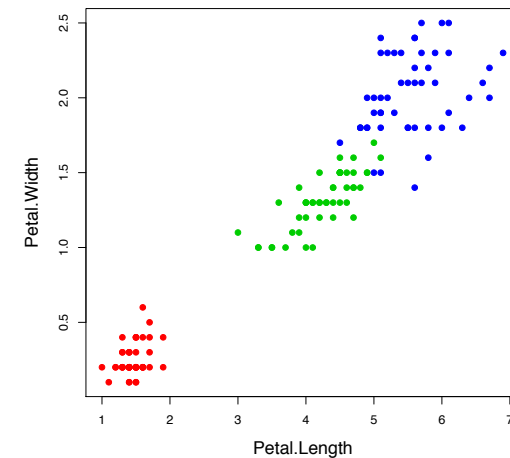


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Iris Dataset

Just focus on two predictor variables.

```
iris.data <- iris[,3:4]
plot(iris.data,col=ct+1,pch=20,cex=1.5,cex.lab=1.4)
```



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Iris Dataset

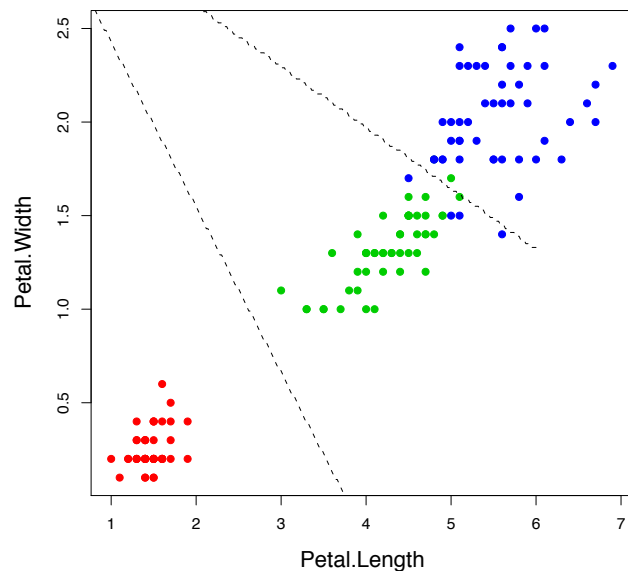
Computing and plotting the LDA boundaries.

```
##fit LDA
iris.lda <- lda(x=iris.data,grouping=ct)

##create a grid for our plotting surface
x <- seq(-6,6,0.02)
y <- seq(-4,4,0.02)
z <- as.matrix(expand.grid(x,y),0)
m <- length(x)
n <- length(y)

##classes are 1,2 and 3, so set contours at 1.5 and 2.5
iris.ldp <- predict(iris.lda,z)$class
contour(x,y,matrix(iris.ldp,m,n),
        levels=c(1.5,2.5), add=TRUE, d=FALSE, lty=2)
```

Iris Dataset



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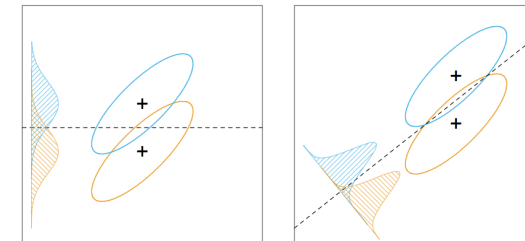
Fisher's Linear Discriminant Analysis

- ▶ In LDA, data vectors are classified based on Mahalanobis distance from cluster means, which lie on a $K - 1$ affine subspace.
- ▶ In measuring these distances, directions orthogonal⁵ to the subspace can be ignored.
- ▶ Projecting data vectors onto the subspace can be viewed as a dimensionality reduction technique that preserves discriminative information about $(y_i)_{i=1}^n$.
- ▶ As with PCA, we can visualize the structure in the data by choosing an appropriate basis for the subspace and projecting data onto it.
- ▶ Choose a basis by finding directions that are separate classes best.

⁵Orthogonality defined in terms of the inner product corresponding to Mahalanobis distance:
 $\langle x, y \rangle = x^T \Sigma^{-1} y$.

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Fisher's Linear Discriminant Analysis



- ▶ Find a direction $v \in \mathbb{R}^p$ to maximize the variance ratio

$$\frac{v^T B v}{v^T \Sigma v}$$

where

$$\Sigma = \frac{1}{n-1} \sum_{i=1}^n (x_i - \mu_{y_i})(x_i - \mu_{y_i})^T \quad (\text{within class covariance})$$

$$B = \frac{1}{n-1} \sum_{k=1}^K n_k (\mu_{y_k} - \bar{x})(\mu_{y_k} - \bar{x})^T \quad (\text{between class covariance})$$

B has rank at most $K - 1$.

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Figure from Hastie et al.

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Discriminant Coordinates

- ▶ To solve for the optimal v , we first reparameterize it as $u = \Sigma^{\frac{1}{2}}v$.

$$\frac{v^T B v}{v^T \Sigma v} = \frac{u^T (\Sigma^{-\frac{1}{2}})^T B \Sigma^{-\frac{1}{2}} u}{u^T u} = \frac{u^T B^* u}{u^T u}$$

where $B^* = (\Sigma^{-\frac{1}{2}})^T B \Sigma^{-\frac{1}{2}}$.

- ▶ The maximization over u is achieved by the first eigenvector u_1 of B^* .
- ▶ We also look at the remaining eigenvectors u_i associated to the non-zero eigenvalues and defined the **discriminant coordinates** as $v_i = \Sigma^{-\frac{1}{2}}u_i$.
- ▶ The v_i 's span exactly the affine subspace spanned by $(\Sigma^{-1}\mu_k)_{k=1}^K$ (these vectors are given as the "linear discriminants" in the R-function `lda`).

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Crabs Dataset

```
library(MASS)
data(crabs)

## numeric and text class labels
ct <- as.numeric(crabs[,1])-1+2*(as.numeric(crabs[,2])-1)

## Projection on Fisher's linear discriminant directions
print(cb.lda <- lda(log(crabs[,4:8]),ct))
```

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Crabs Dataset

```
> > > > > > > Call:
lda(log(crabs[, 4:8]), ct)

Prior probabilities of groups:
 0  1  2  3
0.25 0.25 0.25 0.25

Group means:
      FL      RW      CL      CW      BD
0 2.564985 2.475174 3.312685 3.462327 2.441351
1 2.852455 2.683831 3.529370 3.649555 2.733273
2 2.672724 2.443774 3.437968 3.578077 2.560806
3 2.787885 2.489921 3.490431 3.589426 2.701580

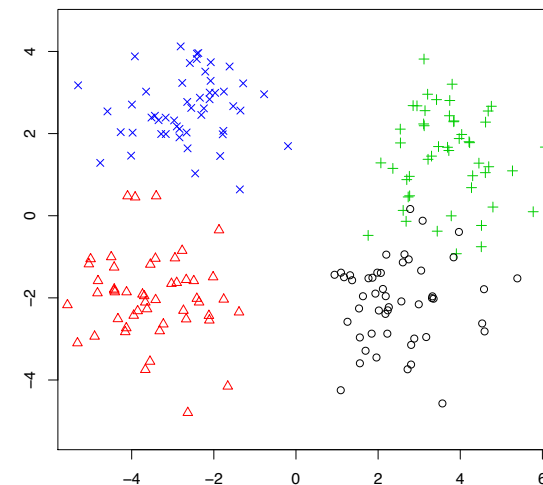
Coefficients of linear discriminants:
      LD1      LD2      LD3
FL -31.217207 -2.851488 25.719750
RW -9.485303 -24.652581 -6.067361
CL -9.822169 38.578804 -31.679288
CW 65.950295 -21.375951 30.600428
BD -17.998493 6.002432 -14.541487

Proportion of trace:
 LD1  LD2  LD3
0.6891 0.3018 0.0091
```

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Crabs Dataset

```
cb.ldap <- predict(cb.lda)
eqsplot(cb.ldap$x, pch=ct+1, col=ct+1)
```



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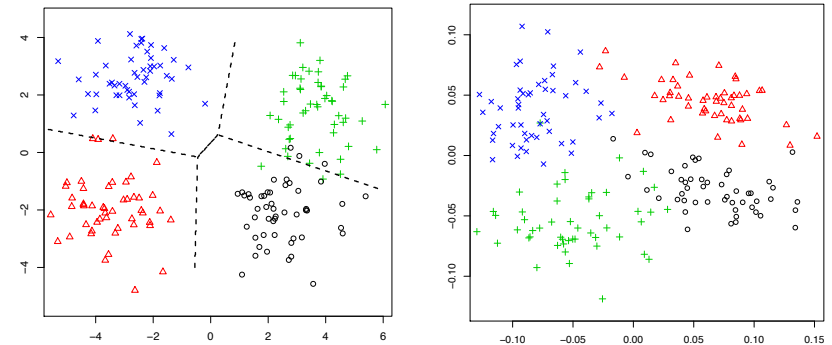
Crabs Dataset

```
## display the decision boundaries
## take a lattice of points in LD-space
x <- seq(-6, 6, 0.02)
y <- seq(-4, 4, 0.02)
z <- as.matrix(expand.grid(x, y, 0))
m <- length(x)
n <- length(y)

## predict onto the grid
cb.ldap <- lda(cb.ldp$x, ct)
cb.ldpp <- predict(cb.ldap, z)$class

## classes are 0,1,2 and 3 so set contours
## at 0.5,1.5 and 2.5
contour(x, y, matrix(cb.ldpp, m, n),
        levels=c(0.5, 1.5, 2.5),
        add=TRUE, d=FALSE, lty=2, lwd=2)
```

Crabs Dataset

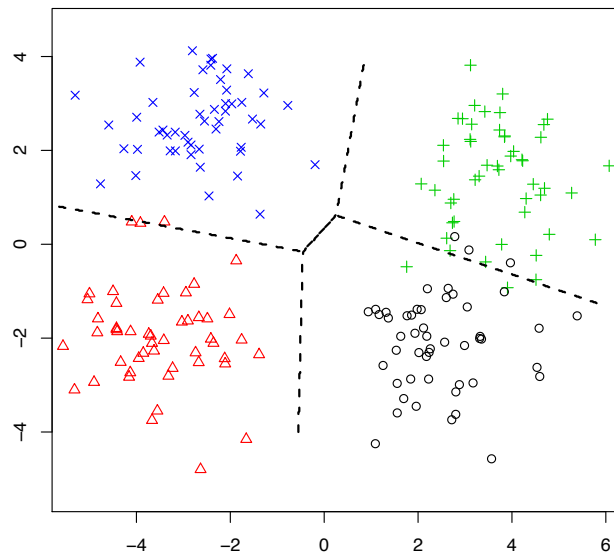


LDA separates the groups better.

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Crabs Dataset



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Naïve Bayes

- ▶ Assume we are interested in classifying documents; e.g. scientific articles or emails.
- ▶ A basic but standard model for text classification consists of considering a pre-specified dictionary of p words (including say physics, calculus.... or dollars, sex etc.) and summarizing each document i by a binary vector x_i where

$$x_{ij} = \begin{cases} 1 & \text{if word } j \text{ is present in document } i \\ 0 & \text{otherwise.} \end{cases}$$

- ▶ To implement a probabilistic classifier, we need to model $f_k(x|\phi_k)$ for each class $k = 1, \dots, K$.

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Naïve Bayes

- ▶ A Naïve Bayes approach ignores feature correlations and assumes $f_k(x) = f(x|\phi_k)$ where

$$f_k(x_i) = f(x_i|\phi_k) = \prod_{j=1}^p (\phi_{kj})^{x_{ij}} (1 - \phi_{kj})^{1-x_{ij}}$$

- ▶ Given dataset, the MLE is easily obtained

$$\hat{\pi}_k = \frac{n_k}{n} \quad \hat{\phi}_{kj} = \frac{\sum_{i:y_i=k} x_{ij}}{n_k}$$

- ▶ One problem: if word j did not appear in documents labelled as class k then $\hat{\phi}_{kj} = 0$ and

$$\mathbb{P}(Y = k | X = x \text{ with } j\text{th entry equal to } 1) = 0$$

i.e. we will never attribute a new document containing word j to class k .

- ▶ This problem is called **overfitting**, and is a major concern in modelling high-dimensional datasets common in machine learning.