Michaelmas Term 2009

- Problems 1
- 1. Which of the following densities are within an exponential family? Explain your reasoning.

(a)

$$f(x,\theta) = (1-\theta)\theta^x; \quad x = 0, 1, 2, \dots$$

where  $0 < \theta < 1$ ;

(b)

$$f(x; \alpha, \lambda) = \frac{\lambda^{\alpha}}{\Gamma(\alpha)} e^{-\lambda x} x^{\alpha - 1}, \quad x > 0,$$

where  $\alpha > 0, \lambda > 0;$ 

(c)

$$f(x,\theta) = e^{-(x-\theta)}, \quad x \ge \theta.$$

- **2.** Suppose  $X_1, X_2, \ldots, X_n$  is a random sample from the Pareto distribution  $f(x, \lambda) = \frac{\lambda \alpha^{\lambda}}{x^{\lambda+1}}$  with  $x > \alpha$ ,  $\lambda > 0$ , and  $\alpha > 0$  known. Find the likelihood function for  $\lambda$ , and find a minimal sufficient statistic for  $\lambda$ .
- **3.** Suppose  $X_1, X_2, \ldots, X_n$  is a random sample from the log-normal distribution with density

$$f(x,\mu,\phi) = \frac{1}{x\sqrt{2\pi\phi}} \exp\left\{-\frac{1}{2\phi}(\log x - \mu)^2\right\}$$

with  $\phi > 0$  (so that  $\log X_j \sim \mathcal{N}(\mu, \phi)$ ). Find a minimal sufficient statistic for the parameter  $\theta = (\mu, \phi)$ .

**4.** Suppose  $X_1, \ldots, X_n$  are independent and exponentially distributed, each with density function

$$f(x;\theta) = \frac{1}{\theta}e^{-x/\theta}, \qquad x \ge 0.$$

Let  $\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$  and put

$$T = \frac{\overline{X}}{\sqrt{\frac{1}{n-1}\sum_{i=1}^{n}(X_i - \overline{X})^2}}$$

Show that T is an ancillary statistic. What does this say about t-tests on exponential data?

**5.** Let  $X_1, \ldots, X_n$  be i.i.d. uniform  $\mathcal{U}\left[\theta - \frac{1}{2}, \theta + \frac{1}{2}\right]$  random variables.

a) Show that  $(X_{(1)}, X_{(n)})$  is minimal sufficient for  $\theta$ .

b) Show that  $(S, A) = (\frac{1}{2}(X_{(1)} + X_{(n)}), X_{(n)} - X_{(1)})$  is minimal sufficient for  $\theta$ , and that the distribution of A is independent of  $\theta$  (so A is an ancillary statistic). c) Show that any value contained in the interval  $[x_{(n)} - \frac{1}{2}, x_{(1)} + \frac{1}{2}]$  is a maximum-likelihood-estimator for  $\theta$ .

- 6. The random variables  $X_1, \ldots, X_n$  are independent with geometric distribution  $\mathsf{P}(X_i = x) = p(1-p)^{x-1}$  for  $x = 1, 2, \ldots$  Let  $\theta = p^{-1}$ .
  - (i) Find the maximum likelihood estimator for p. Considering n = 1, is it unbiased?
  - (ii) Show that  $\hat{\theta} = \overline{X}$  is the maximum likelihood estimator for  $\theta$ . Is it unbiased?
  - (iii) Compute the expected Fisher information for  $\theta$ .
  - (iv) Does  $\hat{\theta}$  attain the Cramer-Rao lower bound?
- 7. Let  $X_1, \ldots, X_n$  be i.i.d.  $U[0, \theta]$ , having density

$$f(x;\theta) = \frac{1}{\theta}, \qquad 0 \le x \le \theta$$

with  $\theta > 0$ .

- (i) Estimate  $\theta$  using both the method of moments and maximum likelihood.
- (ii) Calculate the means and variances of the two estimators.
- (iii) Which one should be preferred and why?
- 8. Suppose  $X_1, \ldots, X_n$  are a random sample with mean  $\mu$  and finite variance  $\sigma^2$ . Use the delta method to show that, in distribution,

$$\sqrt{n}(\overline{X}_n^2 - \mu^2) \to \mathcal{N}(0, 4\mu^2\sigma^2).$$

What would you suggest if  $\mu = 0$ ?