Sparse random graphs with exchangeable point processes

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Joint work with Emily Fox (U. Washington)



Exchangeable matrices and their limitations

Statistical network models using exchangeable random measures

Exchangeability and sparsity properties

Special case: Generalized gamma process

Posterior characterization & Inference

Experimental results

# Outline

#### Introduction

Exchangeable matrices and their limitations

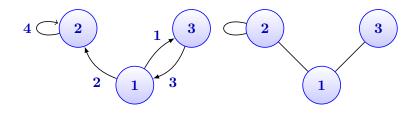
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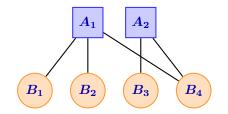
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- Multi-edges directed graphs
  - Emails
  - Citations
  - WWW
- Simple graphs
  - Social network
  - Protein-protein interaction



#### Bipartite graphs

- Scientists authoring papers
- Readers reading books
- Internet users posting messages on forums
- Customers buying items

- Build a statistical model of the network to
  - Find interpretable structure in the network
  - Predict missing edges
  - Predict connections of new nodes

- Properties of real world networks
  - Sparsity

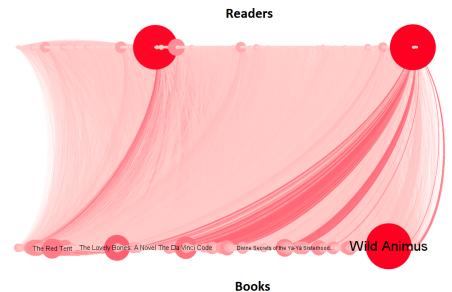
Dense graph:  $n_e = \Theta(n^2)$ Sparse graph:  $n_e = o(n^2)$ 

with  $n_e$  the number of edges and n the number of nodes

Power-law degree distributions

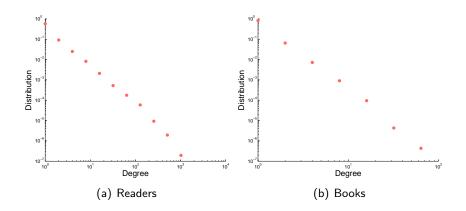
# Book-crossing community network

5 000 readers, 36 000 books, 50 000 edges



# Book-crossing community network

Degree distributions on log-log scale



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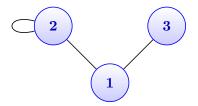
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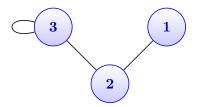
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- Statistical network modeling
- Probabilistic symmetry: exchangeability
- Ordering of the nodes is irrelevant



- Statistical network modeling
- Probabilistic symmetry: exchangeability
- Ordering of the nodes is irrelevant



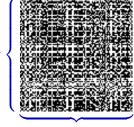
- Graphs usually represented by a discrete structure
- Adjacency matrix  $X_{ij} \in \{0,1\}, (i,j) \in \mathbb{N}^2$

 $\pi$ 

Joint exchangeability

$$(X_{ij}) \stackrel{d}{=} (X_{\pi(i)\pi(j)})$$

for any permutation  $\pi$  of  $\mathbb N$ 



Aldous-Hoover representation theorem

$$(X_{ij})=(F(U_i,U_j,U_{\{ij\}}))$$

where  $U_i, U_{\{ij\}}$  are uniform random variables and F is a random function from  $[0, 1]^3$  to  $\{0, 1\}$ 

 Several network models fit in this framework (e.g. stochastic blockmodel, infinite relational model, etc.)

Corollary of A-H theorem

Exchangeable random graphs are either empty or dense

To quote the survey paper of Orbanz and Roy

"the theory [...] clarifies the limitations of exchangeable models. It shows, for example, that most Bayesian models of network data are inherently misspecified"

 Give up exchangeability for sparsity? e.g. preferential attachment model

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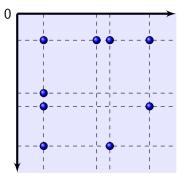
Experimental results

- ▶ Representation of a graph as a (marked) point process over  $\mathbb{R}^2_+$
- Representation theorem by Kallenberg for jointly exchangeable point processes on the plane
- Construction based on a completely random measure
- Properties of the model
  - Exchangeability
  - Sparsity
  - Power-law degree distributions (with exponential cut-off)
  - Interpretable parameters and hyperparameters
  - Reinforced urn process construction
- Posterior characterization
- Scalable inference

 $\blacktriangleright$  Undirected graph represented as a point process on  $\mathbb{R}^2_+$ 

$$Z = \sum_{i,j} z_{ij} \delta_{( heta_i, heta_j)}$$

with  $heta_i \in \mathbb{R}$ ,  $z_{ij} \in \{0,1\}$  with  $z_{ij} = z_{ji}$ 

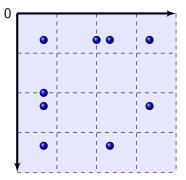


Joint exchangeability

Let  $A_i = [h(i-1), hi]$  for  $i \in \mathbb{N}$  then

$$(Z(A_i imes A_j)) \stackrel{d}{=} (Z(A_{\pi(i)} imes A_{\pi(j)}))$$

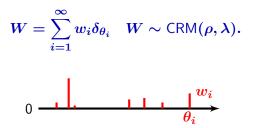
for any permutation  $\pi$  of  $\mathbb N$  and any h>0



- Kallenberg derived a de Finetti style representation theorem for jointly and separately exchangeable point processes on the plane
- Representation via random transformations of unit rate Poisson processes and uniform variables
- Continuous-time equivalent of Aldous-Hoover for binary variables
- Our construction will fit into this framework

## Completely random measures

- Nodes are embedded at some location  $heta_i \in \mathbb{R}_+$
- To each node is associated some sociability parameter  $w_i$
- Homogeneous completely random measure on  $\mathbb{R}_+$



Lévy measure ν(dw, dθ) = ρ(dw)λ(dθ) with λ the Lebesgue measure

[Kingman, 1967] 21 / 57

## Completely random measures

- Lévy measure  $u(dw, d\theta) = \rho(dw)\lambda(d\theta)$  with  $\lambda$  the Lebesgue measure
- ho is a measure on  $\mathbb{R}_+$  such that

$$\int_0^\infty (1 - e^{-w})\rho(dw) < \infty.$$
 (1)

which implies that  $W([0,T]) < \infty$  for any  $T < \infty$ .

 $\int_0^\infty \rho(dw) = \infty \Longrightarrow \text{Infinite number of jumps in any interval } [0,T]$ "Infinite activity CRM"

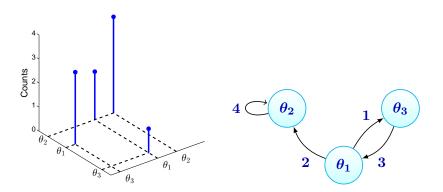
 $\int_0^\infty \rho(dw) < \infty \Longrightarrow \text{Finite number of jumps in any interval } [0,T]$ "Finite activity CRM"

# Model for multi-edges directed graphs

We represent the integer-weighted directed graph using an atomic measure on  $\mathbb{R}^2_+$ 

$$D = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} n_{ij} \delta_{( heta_i, heta_j)},$$

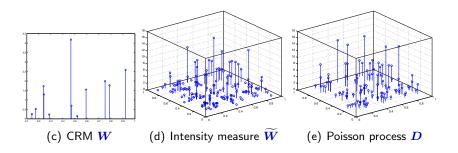
where  $n_{ij}$  counts the number of directed edges from node  $\theta_i$  to node  $\theta_j$ .



# Model for multi-edges directed graphs

► Conditional Poisson process with intensity measure W = W × W on the product space R<sup>2</sup><sub>+</sub>:

 $D \mid W \sim \mathsf{PP}(W \times W)$ 



# Model for multi-edges directed graphs

- ▶ By construction, for any bounded intervals A and B of  $\mathbb{R}_+$ ,  $\widetilde{W}(A \times B) = W(A)W(B) < \infty$
- Finite number of counts over  $A imes B \subset \mathbb{R}^2_+$

 $D(A \times B) < \infty$ 

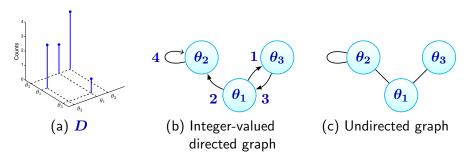
# Model for undirected graphs

Point process

$$Z = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} z_{ij} \delta_{( heta_i, heta_j)},$$

with the convention  $z_{ij} = z_{ji} \in \{0,1\}$ 

• Constructed from D by setting  $z_{ij} = z_{ji} = 1$  if  $n_{ij} + n_{ji} > 0$  and  $z_{ij} = z_{ji} = 0$  otherwise



Model for undirected graphs

Hierarchical model

 $egin{aligned} W &= \sum_{i=1}^\infty w_i \delta_{ heta_i} \ D &= \sum_{ij} n_{ij} \delta_{( heta_i, heta_j)} \ Z &= \sum_{ij} \min(n_{ij}+n_{ji},1) \delta_{( heta_i, heta_j)} \end{aligned}$ 

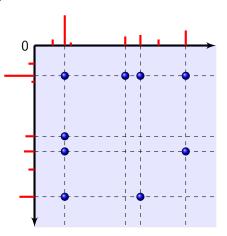
 $W \sim \mathsf{CRM}(
ho, \lambda)$  $D \sim \mathsf{PP}(W imes W)$ 

## Model for undirected graphs

• Equivalent direct formulation for  $i \leq j$ 

$$\Pr(z_{ij} = 1 \mid w) = \left\{egin{array}{cc} 1 - \exp(-2w_iw_j) & i 
eq j \ 1 - \exp(-w_i^2) & i = j \end{array}
ight.$$

and  $z_{ji} = z_{ij}$ 



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### Exchangeability and sparsity properties

Special case: Generalized gamma process

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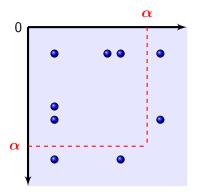
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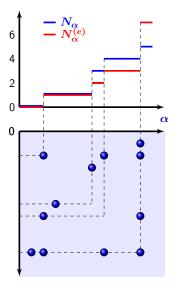
## Properties: Exchangeability

Exchangeability Let h>0 and  $A_i=[h(i-1),hi],\,i\in\mathbb{N}.$  By construction, $(Z(A_i imes A_j))\stackrel{d}{=}(Z(A_{\pi(i)} imes A_{\pi(j)}))$ 

for any permutation  $\pi$  of  $\mathbb{N}$ .

- $W(\mathbb{R}_+) = \infty$ , so infinite number of edges on  $\mathbb{R}^2_+$
- Restrictions  $D_{\alpha}$  and  $Z_{\alpha}$  of D and Z, respectively, to the box  $[0, \alpha]^2$ .
- $N_{lpha}$  number of nodes, and  $N_{lpha}^{(e)}$  number of edges





#### Definition

(Regular variation) Let  $W \sim CRM(\rho, \lambda)$ . The (infinite-activity) CRM is said to be *regularly varying* if the tail Lévy intensity verifies

$$\int_x^\infty 
ho(dw) \stackrel{x\downarrow 0}{\sim} \ell(1/x) x^{-\sigma}$$

for  $\sigma \in (0, 1)$  where  $\ell$  is a slowly varying function satisfying  $\lim_{t\to\infty} \ell(at)/\ell(t) = 1$  for any a > 0.

Assume ho 
eq 0 and  $\mathbb{E}[W([0,1])] < \infty$ .

#### Theorem

Let  $N_{\alpha}$  be the number of nodes and  $N_{\alpha}^{(e)}$  the number of edges in the undirected graph restriction,  $Z_{\alpha}$ . Then

$$N_{\alpha}^{(e)} = \left\{ egin{array}{ll} \Theta\left(N_{lpha}^{2}
ight) & ext{if } W ext{ is finite-activity} \ o\left(N_{lpha}^{2}
ight) & ext{if } W ext{ is infinite-activity} \ O\left(N_{lpha}^{2/(1+\sigma)}
ight) & ext{if } W ext{ is regularly varying}^{1} ext{ with } \sigma \in (0,1) \end{array} 
ight.$$

almost surely as  $\alpha \to \infty$ .

<sup>&</sup>lt;sup>1</sup>with  $\lim_{t\to\infty}\ell(t)>0$ 

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## Generalized Gamma Process

Lévy intensity

$$\frac{1}{\Gamma(1-\sigma)}w^{-1-\sigma}e^{-\tau w}$$

with  $\sigma \in (-\infty, 0]$  and  $\tau > 0$ or  $\sigma \in (0, 1)$  and  $\tau \ge 0$ 

- Special cases:
  - Gamma process ( $\sigma = 0$ )
  - Stable process  $( au = 0, \sigma \in (0, 1))$
  - Inverse Gaussian process  $(\sigma=1/2, au>0)$
- Infinite activity for  $\sigma \geq 0$
- Regularly varying for  $\sigma \in (0, 1)$
- Exact sampling of the graph via an urn process
- Power-law degree distribution

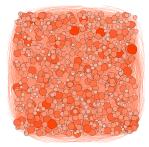
#### Generalized Gamma Process Sparsity

#### Theorem

Let  $N_{\alpha}$  be the number of nodes and  $N_{\alpha}^{(e)}$  the number of edges in the undirected graph restriction,  $Z_{\alpha}$ . Then

$$N_{lpha}^{(e)} = \left\{ egin{array}{ll} \Theta\left(N_{lpha}^2
ight) & ext{if } \sigma < 0 \ o\left(N_{lpha}^2
ight) & ext{if } \sigma \in [0,1), au > 0 \ O\left(N_{lpha}^{2/(1+\sigma)}
ight) & ext{if } \sigma \in (0,1), au > 0 \end{array} 
ight.$$

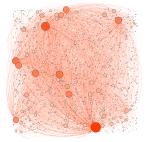
almost surely as  $\alpha \to \infty$ . That is, the underlying graph is sparse if  $\sigma \ge 0$  and dense otherwise.



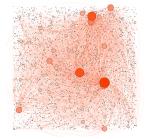
#### Erdös-Rényi G(1000, 0.05)



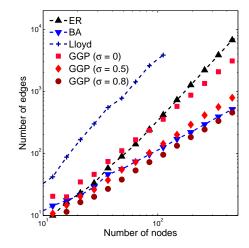
#### Gamma Process



 $\mathsf{GGP}\ (\boldsymbol{\sigma}=\mathbf{0.5})$ 

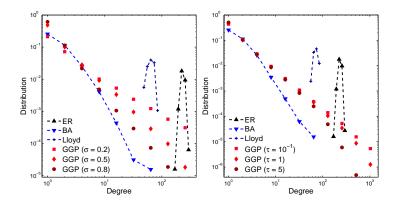


GGP ( $\sigma = 0.8$ )



Power-law degree distributions

- Power-law like behavior providing a heavy-tailed degree distribution
- Higher power-law exponents for larger σ
- The parameter au tunes the exponential cut-off in the tails.



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### Posterior characterization

Conditional distribution of  $W_{\alpha}$  given  $D_{\alpha}$ .

#### Theorem

Let  $(\theta_1, \ldots, \theta_{N_{\alpha}})$ ,  $N_{\alpha} \geq 0$ , be the set of support points of  $D_{\alpha}$  such that  $D_{\alpha} = \sum_{1 \leq i,j \leq N_{\alpha}} n_{ij} \delta_{(\theta_i, \theta_j)}$ . Let  $m_i = \sum_{j=1}^{N_{\alpha}} (n_{ij} + n_{ji}) > 0$  for  $i = 1, \ldots, N_{\alpha}$ . The conditional distribution of  $W_{\alpha}$  given  $D_{\alpha}$  is equivalent to the distribution of

$$w_*\sum_{i=1}^{\infty}\widetilde{P}_i\delta_{\widetilde{ heta}_i}+\sum_{i=1}^{N_{lpha}}w_i\delta_{ heta_i}$$

where  $\tilde{\theta}_i \sim \text{Unif}([0, \alpha])$ , and  $(\tilde{P}_i)|w_* \sim \text{PK}(\rho|w_*)$  are from a Poisson-Kingman distribution. The weights  $(w_1, \ldots, w_{N_{\alpha}}, w_*)$  are jointly dependent conditional on  $D_{\alpha}$ , with  $p(w_1, \ldots, w_{N_{\alpha}}, w_*|D_{\alpha}) \propto$ 

$$\left[\prod_{i=1}^{N_{\alpha}} w_{i}^{m_{i}}\right] e^{-\left(\sum_{i=1}^{N_{\alpha}} w_{i} + \boldsymbol{w}_{\star}\right)^{2}} \left[\prod_{i=1}^{N_{\alpha}} \rho(w_{i})\right] \times g_{\alpha}^{\star}(\boldsymbol{w}_{\star})$$

where  $g_{\alpha}^*$  is the probability density function of the random variable  $W_{\alpha}^* = W_{\alpha}([0, \alpha])$ .

[Prünster, 2002, James, 2002, James et al., 2009] 43/57

## Posterior inference for undirected graphs

- Let  $\phi = (\alpha, \sigma, \tau)$  with flat priors
- We want to approximate

 $p(w_1,\ldots,w_{N_lpha},w_*,\phi|(z_{ij})_{1\leq i,j\leq N_lpha})$ 

- Latent count variables  $\overline{n}_{ij} = n_{ij} + n_{ji}$
- Markov chain Monte Carlo sampler
  - 1. Update the weights  $(w_1, \ldots, w_{N_{\alpha}})$  given the rest using an Hamiltonian Monte Carlo update
  - 2. Update the total mass  $w_*$  and hyperparameters  $\phi = (\alpha, \sigma, \tau)$  given the rest using a Metropolis-Hastings update
  - 3. Update the latent counts  $(\overline{n}_{ij})$  given the rest from a truncated Poisson distribution

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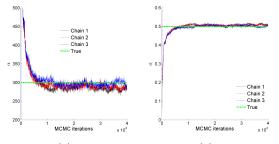
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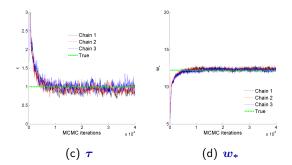
### Simulated data

- $\blacktriangleright$  Simulation of a GGP graph with  $lpha=300, \sigma=1/2, \tau=1$
- 13,995 nodes and 76,605 edges
- MCMC sampler with 3 chains and 40,000 iterations
- Takes 10min on a standard desktop with Matlab



(a) **α** 





### Simulated data

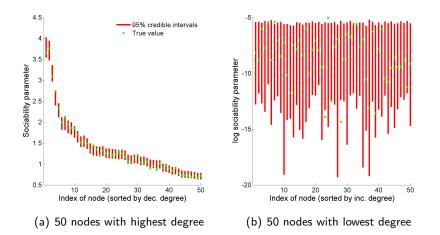
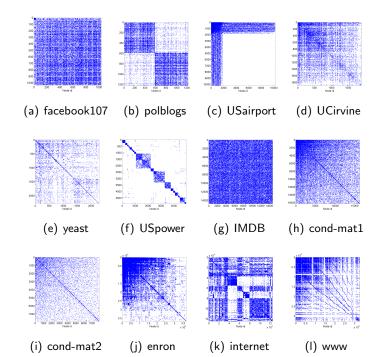


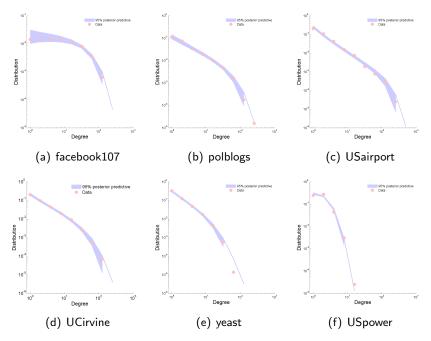
Figure: 95 % posterior intervals of (a) the sociability parameters  $w_i$  of the 50 nodes with highest degree and (b) the log-sociability parameter  $\log w_i$  of the 50 nodes with lowest degree. True values are represented by a green star.

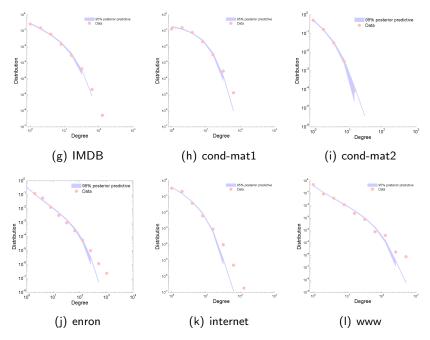
- Assessing the sparsity of the network
- We aim at reporting  $\Pr(\sigma \geq 0|z)$  based on a set of observed connections (z)
- 12 different networks
- $\blacktriangleright \sim 1,000-300,000$  nodes and 10,000-1,000,000 edges



## Real network data

Name	Nb nodes	Nb edges	Time (min)	$\Pr(\sigma > 0 z)$	99% CI <i>σ</i>
facebook107	1,034	26,749	1	0.00	[-1.06, -0.82]
polblogs	1,224	16,715	1	0.00	[-0.35, -0.20]
USairport	1,574	17,215	1	1.00	[0.10, 0.18]
UCirvine	1,899	13,838	1	0.00	[-0.14, -0.02]
yeast	2,284	6,646	1	0.28	$\left[-0.09, 0.05\right]$
USpower	4,941	6,594	1	0.00	[-4.84, -3.19]
IMDB	14,752	38,369	2	0.00	[-0.24, -0.17]
cond-mat1	16,264	47,594	2	0.00	$\left[-0.95,-0.84\right]$
cond-mat2	7,883	8,586	1	0.00	[-0.18, -0.02]
Enron	36,692	183,831	7	1.00	[0.20, 0.22]
internet	124,651	193,620	15	0.00	[-0.20, -0.17]
www	325,729	1,090,108	132	1.00	$\left[\boldsymbol{0.26},\boldsymbol{0.30}\right]$





## Conclusion

- Statistical network models
- Build on exchangeable random measures
- Sparsity and power-law properties
- Scalable inference
- Similar construction for bipartite graphs
- Extensions to more structured models: low-rank, block-model, covariates, dynamic networks,etc

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